Contents lists available at ScienceDirect

European Journal of Operational Research

journal homepage: www.elsevier.com/locate/eor



Decision Support

Optimal insurance design with Lambda-Value-at-Risk

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ARTICLE INFO

Keywords: Risk management Optimal insurance Lambda-Value-at-Risk Model uncertainty Stop-loss

ABSTRACT

This paper explores optimal insurance solutions based on the Lambda-Value-at-Risk (ΛVaR). Using the expected value premium principle, we first analyze a stop-loss indemnity and provide a closed-form expression for the deductible parameter. A necessary and sufficient condition for the existence of a positive and finite deductible is also established. We then generalize the stop-loss indemnity and show that, akin to the VaR model, a limited stop-loss indemnity remains optimal within the AVaR framework. Further, we examine the use of A'VaR as a premium principle and show that full or no insurance is optimal. We also identify that a limited loss indemnity is optimal when Λ' VaR is solely used to determine the risk-loading in the premium principle. Additionally, we investigate the impact of model uncertainty, particularly in scenarios where the loss distribution is unknown but lies within a specified uncertainty set. Our findings suggest that a limited stop-loss indemnity is optimal when the uncertainty set is defined using a likelihood ratio. Meanwhile, when only the first two moments of the loss distribution are available, we provide a closed-form expression for the optimal deductible in a stop-loss indemnity.

1. Introduction

Optimal insurance is the problem of finding the optimal insurance contract. A decision maker (DM) seeks to transfer a portion of his loss to an insurer and will pay a premium for this coverage. The premium is traditionally calculated as the expected value of the insured risk, plus a risk loading. The literature on optimal insurance originated from the early works of Borch (1960) and Arrow (1963), which focused on mean-variance and expected utility preferences, respectively. Many authors have extended these settings by considering alternative objective functions such as behavioral economic preferences and risk measures for evaluating future losses after insurance. For instance, Cai and Tan (2007), Cai et al. (2008), and Chi and Tan (2011) study the optimal insurance problem by minimizing the Value-at-Risk (VaR) and the Conditional Value-at-Risk (CVaR) of the DM's total risk exposure. Subsequently, Cui et al. (2013) and Assa (2015) generalize these settings with distortion risk measures, which include VaR and CVaR as special cases.

In this paper, we explore optimal insurance problems using the Lambda-Value-at-Risk (AVaR) as the DM's risk measure. The AVaR is introduced by Frittelli et al. (2014), and has recently gained academic interest as a generalization of the well-known VaR. Unlike VaR, which uses a fixed confidence level, ΛVaR incorporates a function Λ as its confidence level that varies with the underlying loss realizations. This feature provides greater flexibility, allowing one to adjust Λ to either increase sensitivity to extreme losses or focus more on the most likely outcomes. Similar to VaR, and in contrast to CVaR, AVaR is always finite (except for probability levels 0 and 1), even for heavytailed losses with infinite expectations. While AVaR is a monotonic risk measure, it differs from VaR in that it is not cash- or comonotonic additive. A characterization of AVaR is provided by Bellini and Peri (2022), who formalize a crucial property called locality. This, together with monotonicity, normalization, and weak semi-continuity, uniquely characterizes ΛVaR with a non-increasing function $\Lambda(x)$ on \mathbb{R} . AVaR has also been studied from various perspectives, including robustness, elicitability and consistency (see Burzoni et al., 2017), estimation and backtesting (see Corbetta & Peri, 2018; Hitaj et al., 2018), quasi-convexity, cash subadditivity and quasi-star-shapedness (see Han et al., 2025), robust formulations (see Han & Liu, 2024), and capital allocation with the Euler rule (see Ince et al., 2022).

An optimal insurance problem based on AVaR with the expected value premium is first considered by Balbás et al. (2023), as an application to their main theory on Λ -fixed points. Their result assumes finite support of the function Λ and solves the optimization problem

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using a linear programming method. As a significant improvement to the existing discussion, we solve the ΛVaR -based optimal insurance problem with general decreasing Λ functions and loss distributions by more concise theoretical proofs. Moreover, various assumptions of indemnities and premium principles are considered. To the best of our knowledge, this work is the first to comprehensively study the optimal (re)insurance design problem with ΛVaR in a general setup.

We first focus on the problem with the expected premium principle. Assuming that the admissible set of indemnities is of stop-loss types, we present a closed-form expression for the optimal deductible. Moreover, we identify a necessary and sufficient condition for the existence of a positive and finite optimal deductible. Compared with the corresponding condition in the VaR-based optimal insurance problem of Cai and Tan (2007), our condition is less stringent. Next, we consider the more general class of incentive-compatible indemnity functions (Huberman et al., 1983), denoted by \mathcal{F} . Within the class \mathcal{F} , Chi and Tan (2011) show that a limited stop-loss indemnity function is optimal for a VaR-based optimal insurance problem. We extend this finding to Λ VaR, with which the DM can adjust the upper limit of the ceded loss function based on various model parameters and insurance premiums.

Following this, we proceed to consider different premium principles with indemnities in \mathcal{F} . If a Λ' VaR premium principle is applied, we demonstrate that full or zero insurance is the optimal indemnity, with the optimal structure determined solely by the ordering of the Λ VaR and Λ' VaR of the total loss. In the case of a more general Λ' VaR-based premium principle, where Λ' VaR is used to determine a risk loading on top of the expected value of losses, we show that a limited loss indemnity is optimal.

The impact of model uncertainty in optimization based on risk measures has been increasingly recognized in academic research, leading to extensive studies on the topic. For instance, distributionally robust portfolio optimization has been explored by Blanchet et al. (2022), Glasserman and Xu (2013), Huang et al. (2010), Natarajan et al. (2010), Pesenti et al. (2024) and Shao and Zhang (2023). More recently, model uncertainty has also been extensively studied in the context of insurance and reinsurance. Asimit et al. (2017) study the robust and Pareto optimality of insurance contracts; Liu and Mao (2022) propose a robust reinsurance strategy under moment constraints; Birghila et al. (2023) and Boonen and Jiang (2024) examine robust insurance problems by assuming that the uncertainty set is defined using Wasserstein distance; Cai et al. (2024) study the stop-loss and limited loss strategies when the uncertainty set is characterized by mean, variance, and Wasserstein distance; Landriault et al. (2024) propose a noncooperative optimal reinsurance problem incorporating likelihood ratio uncertainty.

This paper presents an analysis of the impact of model uncertainty on the AVaR optimal insurance problem. We assume that the underlying loss distribution is unknown, but it is an element of a so-called uncertainty set, which can be thought of as a collection of plausible loss distribution functions. If such a set is generated by a ball centered around the probability measure used for pricing, measured using a likelihood ratio, then a limited stop-loss indemnity is shown to be optimal within \mathcal{F} . When uncertainty is characterized by knowledge of the first two moments of the loss variable, and we restrict our focus to stoploss indemnities, a closed-form expression for the optimal deductible is provided. In general, we find that when the uncertainty set is defined by any collection of loss distributions, the robust AVaR optimization problem can be transformed into a robust VaR optimization problem over the same uncertainty set. Subsequently, the optimal probability level for AVaR can be determined based on the results from the VaR problem. This observation is consistent with Han and Liu (2024), which does not incorporate an insurance strategy.

The findings of this paper are valuable to both academics and practitioners, as they provide insights into the optimal indemnities associated with a new risk measure, extending beyond the well-studied class of distortion risk measures. Similar to distortion risk measures,

we show that optimal indemnities typically exhibit a piecewise-linear structure, with proportional or co-insurance generally being suboptimal. Specifically, our results complement the existing literature on optimal (re)insurance with risk measures, including but not limited to Asimit et al. (2017), Asimit and Boonen (2018), Assa et al. (2021), Boonen and Jiang (2024), and Tan et al. (2020).

We note that the optimal insurance problems considered in this paper are different from the AVaR-based risk sharing problems in the literature (e.g., Liu, 2024; Liu et al., 2024; Xia & Hu, 2024). In optimal risk sharing problems, the aim is to determine an optimal set of random variables that aggregate to the total risk in the market, while minimizing the sum of risk measures. This objective differs from that of optimal insurance problems, where the goal is to minimize a risk measure for a single agent. Moreover, in optimal insurance problems, the admissible set of indemnities is typically constrained to a specific class of non-decreasing 1-Lipschitz functions, as discussed in Assa (2015). This assumption implies that the retained and ceded losses in optimal insurance problems are comonotonic. In sharp contrast, optimal risk sharing allocations are not comonotonic as shown by Liu (2024) and Liu et al. (2024). Additionally, in optimal insurance, the insurer's objectives are summarized through a premium principle, which may include a risk loading.

The paper is organized as follows. Section 2 introduces ΛVaR and the optimal insurance problem. Section 3 examines optimal insurance indemnities under the expected value premium principle. Section 4 investigates optimal insurance when the premium principle incorporates $\Lambda' \text{VaR}$. Section 5 discusses optimal insurance indemnities under model uncertainty. Section 6 presents some numerical examples, and Section 7 concludes.

2. Problem formulation

Let $(\Omega,\mathcal{G},\mathbb{P})$ be a probability space. Each random variable represents a random risk that is realized at a well-defined future period. Throughout the paper, "increasing" and "decreasing" are in the nonstrict (weak) sense. Let \mathcal{X} be a convex cone of random variables. For any $Z \in \mathcal{X}$, the cumulative distribution function (CDF) associated with Z is denoted by F_Z . For a function $f: \mathbb{R} \to \mathbb{R}$ and $x \in \mathbb{R}$, we write $\lim_{t \uparrow x} f(t) = f(x-)$.

A random vector (Z_1,\ldots,Z_n) is comonotonic if there exists a random variable Z and increasing functions f_1,\ldots,f_n on $\mathbb R$ such that $Z_i=f_i(Z)$ a.s. for every $i=1,\ldots,n$. We define the following properties for a mapping $\rho:\mathcal X\to\mathbb R$:

- [A1] (Monotonicity) $\rho(Y) \leqslant \rho(Z)$ for all $Y, Z \in \mathcal{X}$ with $Y \leqslant Z$;
- [A2] (Cash additivity) $\rho(Y + c) = \rho(Y) + c$ for all $c \in \mathbb{R}$;
- [A3] (Comonotonic additivity) $\rho(Y+Z) = \rho(Y) + \rho(Z)$ whenever Y and Z are comonotonic.

Using the standard terminology in Föllmer and Schied (2016), a mapping $\rho: \mathcal{X} \to \mathbb{R}$ is called a *monetary risk measure* if it satisfies [A1] and [A2].

In this paper, we consider an application of ΛVaR in optimal insurance design problems. Suppose that a decision maker (DM) faces a non-negative random loss $X \in \mathcal{X}_+$, where

$$\mathcal{X}_+ = \{X \in \mathcal{X} \mid X \geqslant 0, \ \mathbb{P}(X < \infty) = 1\}.$$

For generality, we do not assume that X has a finite expectation as infinite-mean models are particularly useful in modeling large insurance claims and catastrophic losses; see, e.g., Chen and Wang (2025). For $\Lambda: \mathbb{R}_+ \to [0,1]$, the risk measure *Lambda-Value-at-Risk* (Λ VaR, see Frittelli et al., 2014) is defined as

$$\Lambda \text{VaR}(X) = \inf\{x \in \mathbb{R}_+ : \mathbb{P}(X \leqslant x) \geqslant \Lambda(x)\}, \quad X \in \mathcal{X}_+. \tag{1}$$

If $\Lambda = \alpha \in [0, 1]$, then Λ VaR boils down to the *Value-at-Risk (VaR)* at level α given by

$$\operatorname{VaR}_{\alpha}(X) = \inf \left\{ x \in \mathbb{R}_{+} : \mathbb{P}(X \leqslant x) \geqslant \alpha \right\}, \quad X \in \mathcal{X}_{+}.$$

By convention, we use $\inf \emptyset = \infty$ and $\sup \emptyset = 0$. It is well known that VaR satisfies all three properties [A1]–[A3]. However, Λ VaR satisfies [A1] but generally does not satisfy [A2] or [A3], and thus does not belong to the class of monetary risk measures. For more properties of Λ VaR, we refer to Bellini and Peri (2022), Han and Liu (2024) and Xia and Hu (2024).

We focus on a special class of functions \mathcal{D} defined below

$$\mathcal{D} = \{ \Lambda : \mathbb{R}_+ \to [0,1] \mid \Lambda \text{ is decreasing with } 0 < \Lambda(0) < 1 \}.$$

In this setup, higher losses are tolerated with a higher probability by the DM. Although (Frittelli et al., 2014) mainly study increasing Λ , the recent works of Burzoni et al. (2017) and Bellini and Peri (2022) have shown that using a decreasing Λ leads to many advantages, including robustness, elicitability, and an axiomatic characterization. In addition, Λ VaR with decreasing Λ functions also satisfy cash-subadditivity and quasi-star-shapedness; see El Karoui and Ravanelli (2009) and Han et al. (2025). For all these reasons, we assume that Λ is a decreasing function in this paper.

Under an insurance contract, the insurer agrees to cover part of the loss X and requires a premium in return. The function $f: \mathbb{R}_+ \to \mathbb{R}_+$ is commonly described as the indemnity or the ceded loss function, while $R(x) \triangleq x - f(x)$ is known as the retained loss function. To prevent the potential ex post moral hazard, where the DM might be incentivized to manipulate the losses, we follow the literature, such as Huberman et al. (1983) and Carlier and Dana (2003), and impose the incentive compatibility condition on the indemnity functions. That is, we consider insurance contracts $f \in \mathcal{F}$, where

$$\mathcal{F} := \left\{ f : \mathbb{R}_+ \to \mathbb{R}_+ \mid f(0) = 0 \text{ and } 0 \leqslant f(x) - f(y) \leqslant x - y, \right.$$
for all $0 \leqslant y \leqslant x$. (2)

Obviously, for any $f \in \mathcal{F}$, f(x) and x - f(x) are increasing in x, and any $f \in \mathcal{F}$ is 1-Lipschitz continuous. The assumption that $f \in \mathcal{F}$ is common in the literature; see e.g., Assa (2015) and the review paper by Cai and Chi (2020). This class is rich enough and includes many commonly used indemnity functions, such as the stop-loss function $f(x) = (x - d)_+$ for some $d \ge 0$ and the quota-share function f(x) = qx for some $q \in [0, 1]$. Since Lipschitz-continuous functions are absolutely continuous, they are almost everywhere differentiable. It follows directly that f can be written as the integral of its derivatives essentially bounded by 1, and $\mathcal F$ can be represented as

$$\mathcal{F} = \left\{ f : \mathbb{R}_+ \to \mathbb{R}_+ \mid f(x) = \int_0^x q(t) \, \mathrm{d}t, \ 0 \le q \le 1, \ x \in \mathbb{R}_+ \right\},\tag{3}$$

where q is called the marginal indemnity function (Assa, 2015; Zhuang et al., 2016).

For a given $f\in\mathcal{F}$, the insurer prices ceded losses using an exogenously given premium principle $\Pi:\mathcal{X}\mapsto\mathbb{R}_+$. The premium is thus given by $\Pi(f(X))$, which is non-negative. The risk exposure of the DM after purchasing insurance is given by

$$T_f(X) = X - f(X) + \Pi(f(X)).$$
 (4)

We consider the following optimization problem:

$$\inf_{f \in \mathcal{F}} \Lambda \text{VaR}(T_f(X)), \tag{5}$$

where we aim to find an optimal ceded loss function such that the risk of the DM is minimized when evaluated by Λ VaR. The following lemma is useful to solve the optimization problem (5).

Lemma 1 (Theorem 3.1 of Han et al., 2025). For $X \in \mathcal{X}_+$ and $\Lambda \in \mathcal{D}$, the risk measure ΛVaR in (1) has the representation

$$\Lambda \operatorname{VaR}(X) = \inf_{x \in \mathbb{R}_+} \left\{ \operatorname{VaR}_{\Lambda(x)}(X) \vee x \right\} = \sup_{x \in \mathbb{R}_+} \left\{ \operatorname{VaR}_{\Lambda(x)}(X) \wedge x \right\}. \tag{6}$$

In fact, (6) holds for any $X \in \mathcal{X}$ and any decreasing $\Lambda : \mathbb{R} \to [0,1]$ that is not constantly 0.

Remark 1. Note that Lemma 1 does not hold for a general Λ even when Λ is increasing. For example, consider $\Lambda(x)=(1+x)/4$ for $0 \le x < 1$, and $\Lambda(x)=\frac{1}{2}$ for $x \ge 1$. Let X be a random variable with $X \sim U(0,1)$. We can compute $\Lambda \operatorname{VaR}(X)=1/3 > \inf_{x \in \mathbb{R}} \left\{ \operatorname{VaR}_{\Lambda(x)}(X) \vee x \right\} = 1/4$, showing that the infimum representation does not hold.

Using Lemma 1, we can transform the optimization problem in (5) into an optimization problem based on VaR.

Proposition 1. For $X \in \mathcal{X}_+$ and $\Lambda \in \mathcal{D}$, we have

$$\inf_{f \in \mathcal{F}} \Lambda \operatorname{VaR}(T_f(X)) = \inf_{x \in \mathbb{R}_+} \left\{ \inf_{f \in \mathcal{F}} \left\{ \operatorname{VaR}_{\Lambda(x)}(X) - \operatorname{VaR}_{\Lambda(x)}(f(X)) + \Pi(f(X)) \right\} \vee x \right\}. \tag{7}$$

Proof. By Lemma 1, we have

$$\begin{split} \inf_{f \in \mathcal{F}} \Lambda \mathrm{VaR}(T_f(X)) &= \inf_{f \in \mathcal{F}} \inf_{x \in \mathbb{R}_+} \left\{ \mathrm{VaR}_{\Lambda(x)}(T_f(X)) \vee x \right\} \\ &= \inf_{x \in \mathbb{R}_+} \inf_{f \in \mathcal{F}} \left\{ \mathrm{VaR}_{\Lambda(x)}(T_f(X)) \vee x \right\} \\ &= \inf_{x \in \mathbb{R}_+} \left\{ \begin{array}{l} \inf_{f \in \mathcal{F}} \left\{ \mathrm{VaR}_{\Lambda(x)}(X) - \mathrm{VaR}_{\Lambda(x)}(f(X)) + H(f(X)) \right\} \vee x \end{array} \right\}, \end{split}$$

where the last equality is because VaR is cash additive and comonotonic additive, and the term x is independent of $f \in \mathcal{F}$. The exchange of the two infima is valid because the minimization over f and x are independent. The proof is complete. \square

Proposition 1 shows that, under certain assumptions on Λ , the optimization problem in (5) can be decomposed into two steps:

- (i) For a fixed x (i.e., a fixed confidence level $\Lambda(x)$), we solve for the optimal ceded insurance contract $f(X) = f(X, \Lambda(x))$ that minimizes $\operatorname{VaR}_{\Lambda(x)} \left(T_f(X) \right)$.
- (ii) We determine the optimal x^* by solving $\inf_{x \in \mathbb{R}_+} \left\{ \operatorname{VaR}_{A(x)} \left(T_{f(X,A(x))}(X) \right) \vee x \right\}$. The optimal ceded function is then given by $f^*(X) = f(X,A(x^*))$.

3. Optimal results under expected value premium principle

In this section, we assume that the insurer prices the indemnity functions using the expected value premium principle, a method that is widely adopted in insurance literature due to its simplicity and economic relevance:

$$\Pi(f(X)) = (1 + \theta)\mathbb{E}[f(X)],\tag{8}$$

where $\theta \geqslant 0$ represents the safety loading parameter. Here, this premium principle can be seen as the indifference premium of a risk-neutral insurer, and θ reflects any additional costs. That is, it is a solution of the following equation:

$$\mathbb{E}[-W + f(X) - \Pi(f(X)) + C(f(X))] = \mathbb{E}[-W], \tag{9}$$

where $W \in \mathbb{R}$ is the initial wealth, and $C(f(X)) = \theta \mathbb{E}[f(X)]$ represents the costs of the insurer, assumed to be linear in the expected indemnity.

3.1. Optimal retention with deductible contract

We first focus on a specific but rather important class of insurance contracts in \mathcal{F} , the stop-loss insurance contracts. Within this class, the insurer covers losses on the DM above a deductible $l \in [0, \infty]$, that

¹ For K > 0, a function $f: \mathbb{R} \to \mathbb{R}$ is K-Lipschitz continuous if |f(x) - f(y)| ≤ K|x - y| for all $x, y \in \mathbb{R}$. We call K the Lipschitz constant.

is, $f(x) = (x - l)_+$ for $x \in \mathbb{R}_+$. We are particularly interested in stoploss contracts because they are widely applied in practice, especially in the property and casualty insurance sectors. Moreover, they have been extensively studied in the literature; see, e.g. Borch (1960), Cai et al. (2024), Cheung et al. (2014), Klages-Mundt and Minca (2020), and Liu and Mao (2022). The risk exposure of the DM after purchasing insurance is given by

$$T_l(X) = X \wedge l + (1 + \theta)\mathbb{E}[(X - l)_+].$$
 (10)

Next, we address the following optimization problem:

$$\inf_{l \in \mathbb{R}} \Lambda \text{VaR}(T_l(X)) \tag{11}$$

by sequentially solving the two optimization problems in (7); note that Proposition 1 also holds if $\mathcal F$ is changed to a different admissible set of indemnities. In particular, the inner optimization problem is addressed by Theorem 2.1 in Cai and Tan (2007). However, as we relax the constraints on the loss distribution, specifically the assumption that the distribution function is continuous and strictly increasing, a modified proof is provided to accommodate this relaxation.

Let $d^*=\mathrm{VaR}_{\theta^*}(X),$ where $\theta^*=\theta/(1+\theta).$ It is clear that $d^*=0$ for $\theta=0.$ Write

$$M = d^* + (1 + \theta)\mathbb{E}[(X - d^*)_+].$$

Theorem 1. Let $\Lambda \in \mathcal{D}$. The solution $l^* \in [0, \infty]$ that solves the $\Lambda \text{VaR-based insurance model } (11)$ is given by

$$l^* = \begin{cases} d^*, & \text{if } M < \Lambda \text{VaR}(X), \\ \infty, & \text{otherwise} \end{cases}$$
 (12)

and the corresponding minimum is given by

 $\Lambda VaR(T_{l^*}(X)) = M \wedge \Lambda VaR(X).$

Proof. If $\mathbb{E}[X] = \infty$, then $M = \infty$. As $\mathbb{P}(X < \infty) = 1$, we have $\Lambda \text{VaR}(X) < \Lambda \text{VaR}(T_l(X)) = \infty$ for all $l \in [0,\infty)$. Therefore, $l^* = \infty$ and $\Lambda \text{VaR}(T_{l^*}(X)) = \Lambda \text{VaR}(X) < \infty = M$. For the rest of the proof, we assume that $\mathbb{E}[X] < \infty$. By Proposition 1, we have

$$\inf_{l \in \mathbb{R}_{+}} \Lambda \operatorname{VaR}(T_{l}(X)) = \inf_{x \in \mathbb{R}_{+}} \left\{ \inf_{l \in \mathbb{R}_{+}} \left\{ \operatorname{VaR}_{\Lambda(x)}(X \wedge l) + (1 + \theta) \mathbb{E}[(X - l)_{+}] \right\} \vee x \right\}.$$
(13)

For a fixed $x \in \mathbb{R}_+$, observe from (13) that

$$g(l) := \text{VaR}_{A(x)}(X \wedge l) + (1 + \theta)\mathbb{E}[(X - l)_{+}]$$

is decreasing for $l \in [\operatorname{VaR}_{A(x)}(X), \infty)$, with the limit $\operatorname{VaR}_{A(x)}(X)$ as $l \to \infty$. For $l \in [0, \operatorname{VaR}_{A(x)}(X))$, g(l) decreases for $l \in [0, \operatorname{VaR}_{A(x)}(X) \wedge d^*)$ and increases for $l \in [\operatorname{VaR}_{A(x)}(X) \wedge d^*, \operatorname{VaR}_{A(x)}(X))$. Therefore, the function g(l) attains its minimum value at either $d^* \geqslant 0$ or as $l \to \infty$.

When $VaR_{A(x)}(X) > M$, we have

$$\inf_{l\in\mathbb{R}_+}\left\{\mathrm{VaR}_{A(x)}(X\wedge l)+(1+\theta)\mathbb{E}[(X-l)_+]\right\}=M,$$

with the corresponding optimal retention given by $l^* = d^*$; when $\operatorname{VaR}_{A(x)}(X) \leqslant M$, we have

$$\inf_{l\in\mathbb{R}_+}\left\{\mathrm{VaR}_{A(x)}(X\wedge l)+(1+\theta)\mathbb{E}[(X-l)_+]\right\}=\mathrm{VaR}_{A(x)}(X),$$

with the corresponding optimal retention given by $l^* = \infty$. For $x \in \mathbb{R}_+$, define

$$M(x) = M \mathbb{1}_{\{ \text{VaR}_{A(x)}(X) > M \}} + \text{VaR}_{A(x)}(X) \mathbb{1}_{\{ \text{VaR}_{A(x)}(X) \le M \}}.$$

Hence, (13) becomes

$$\inf_{l>0} \Lambda \mathrm{VaR}(T_l(X)) = \inf_{x \in \mathbb{R}_+} \left\{ M(x) \vee x \right\}.$$

Let

$$\bar{x} = \sup\{x \in \mathbb{R}_+ : \operatorname{VaR}_{\Lambda(x)}(X) > M\} \text{ and } x^* = \inf\{x \in \mathbb{R}_+ : M(x) \leqslant x\}.$$

Here, \bar{x} can take the value ∞ , but we have $x^* < \infty$. Therefore, if $x^* > \bar{x}$, as Λ is decreasing, we have

$$\inf_{l>0} \Lambda \operatorname{VaR}(T_l(X)) = \inf_{x \in \mathbb{R}_+} \left\{ M(x) \vee x \right\} = \inf_{x \in \mathbb{R}_+} \left\{ \operatorname{VaR}_{\Lambda(x)}(X) \vee x \right\}$$
$$= \Lambda \operatorname{VaR}(X).$$

Moreover, we have

$$\inf_{l\in\mathbb{R}_+} \Lambda \mathrm{VaR}(T_l(X)) = \inf_{x\in\mathbb{R}_+} \left\{ M(x) \vee x \right\} \leqslant \inf_{x\in\mathbb{R}_+} \left\{ M \vee x \right\} = M.$$

Thus, if $x^* > \bar{x}$, $\inf_{l \in \mathbb{R}_+} \Lambda \text{VaR}(T_l(X)) = \Lambda \text{VaR}(X) \leqslant M$. If $x^* \leqslant \bar{x}$, $\inf_{l \in \mathbb{R}_+} \Lambda \text{VaR}(T_l(X)) = M < \Lambda \text{VaR}(X)$. The inequality holds as

$$\begin{split} \Lambda \mathrm{VaR}(X) &= \inf_{x \in \mathbb{R}_+} \left\{ \mathrm{VaR}_{\Lambda(x)}(X) \vee x \right\} > \inf_{x \in \mathbb{R}_+} \left\{ M(x) \vee x \right\} \\ &= \inf_{l \in \mathbb{R}_+} \Lambda \mathrm{VaR}(T_l(X)). \end{split}$$

Thus we have the desired result. \square

Remark 2. When a quota-share insurance contract is employed, the optimal control takes a 'bang–bang' type, with the values determined by the relationship between $\operatorname{VaR}_\alpha(X)$ and $(1+\theta)\mathbb{E}[X]$. This characteristic arises due to the positive homogeneity of VaR. Given that the problem of minimizing $\Lambda \operatorname{VaR}$ can be transformed into a VaR minimization problem, following the steps similar to those in the proof of Theorem 1, it is not difficult to verify that the optimal solution for $\Lambda \operatorname{VaR}$ under the quota-share insurance contract also retains the 'bang–bang' form.

We can see from Theorem 1 that the deductible in (12) coincides with that for minimizing $\operatorname{VaR}_{\alpha}(T_l)$ with $\alpha \in (0,1)$ when $M \leq \min(\operatorname{VaR}_{\alpha}(X), \Lambda \operatorname{VaR}(X))$. When $l^* = \infty$, it indicates that the DM buys no insurance. The formulation of l^* depends solely on the safety loading factor and remains entirely unaffected by the functional form of Λ . Moreover, we have $f^*(x) = x$ when $\theta = 0$, which implies that the DM chooses to transfer all the risk to the insurer when $\mathbb{E}[X] \leq \Lambda \operatorname{VaR}(X)$ (see Eq. (12)) as the premium represents a "cheap" insurance.

In the next result, we provide a necessary and sufficient condition for the existence of a positive and finite deductible, indicating that the DM always chooses to transfer part of the claim to the insurer.

Theorem 2. Assume that $\mathbb{E}[X] < \infty$ and $\Lambda \in D$. The solution $0 < l^* < \infty$ for problem (11) exists if and only if $\theta^* > F_X(0)$ and $\Lambda(M - \epsilon) > F_X(M - \epsilon)$ for all $\epsilon \in (0, M]$.

Proof. We first show the "if" part. If $\theta^* > F_X(0)$, we have $\text{VaR}_{\theta^*}(X) > 0$. By the proof of Theorem 1, we can write

$$\inf_{l \in \mathbb{R}_+} \Lambda \mathrm{VaR}(T_l(X)) = \inf_{\mathrm{VaR}_{\Lambda(x)}(X) \geqslant M} \left\{ M \vee x \right\} \wedge \inf_{\mathrm{VaR}_{\Lambda(x)}(X) < M} \left\{ \mathrm{VaR}_{\Lambda(x)}(X) \vee x \right\}.$$

For $x \in \mathbb{R}_+$ such that $\operatorname{VaR}_{\Lambda(x)}(X) \geqslant M$, we have $\Lambda(x) \geqslant F_X(M)$. Because Λ is decreasing, we have $x \in [0, a]$ for some $a \geqslant 0$ and thus

$$\inf_{\operatorname{VaR}_{A(x)}(X)\geqslant M} \{M\vee x\} = M.$$

For $x \in \mathbb{R}_+$ such that $\operatorname{VaR}_{\Lambda(x)}(X) < M$, we have $\Lambda(x) \leqslant F_X(M)$. Since $\Lambda(M-\epsilon) > F_X(M-\epsilon)$ for all $\epsilon \in (0,M]$ and F_X is continuous, we have $\Lambda(M-) \geqslant F_X(M) \geqslant \Lambda(x)$. When $\Lambda(M-) > F_X(M)$, we have $\Lambda(M-) > \Lambda(x)$ and thus $x \geqslant M > \operatorname{VaR}_{\Lambda(x)}(X)$. It follows that

$$\inf_{\text{VaR}_{A(x)}(X) < M} \left\{ \text{VaR}_{A(x)}(X) \lor x \right\} = \inf_{\text{VaR}_{A(x)}(X) < M} x \geqslant M. \tag{14}$$

When $\varLambda(M-)=F_X(M)$, we have $\mathrm{VaR}_{\varLambda(M-)}(X)=M$. Thus $x < M \leqslant \mathrm{VaR}_{\varLambda(x)}(X)$ for all x < M and $x \geqslant \mathrm{VaR}_{\varLambda(x)}(X)$ for all $x \geqslant M$. It follows that

$$\inf_{\operatorname{VaR}_{A(x)}(X) < M} \left\{ \operatorname{VaR}_{A(x)}(X) \vee x \right\} = \inf_{\operatorname{VaR}_{A(x)}(X) < M, \ x \geqslant M} x \geqslant M. \tag{15}$$

Therefore, we have by (14) and (15) that $\inf_{\text{VaR}_{A(x)}(X) < M} \{ \text{VaR}_{A(x)}(X) < X \} \ge M$.

According to the discussion above, we have

$$\inf_{l \in \mathbb{R}_+} \Lambda \operatorname{VaR}(T_l(X)) = \inf_{\operatorname{VaR}_{A(X)}(X) \geqslant M} \{ M \vee x \} = M,$$

and the corresponding optimal retention is given by $l^* = VaR_{\theta^*}(X)$.

To prove the "only if" part for the existence of optimal retention, we first suppose for contradiction that $\theta^* \leqslant F_X(0)$. It follows that $l^* = \operatorname{VaR}_{\theta^*}(X) \leqslant 0$. This leads to a contradiction with the fact that $l^* > 0$. Next, we suppose for contradiction that $\Lambda(b) \leqslant F_X(b)$ for some b < M. It follows that $\operatorname{VaR}_{\Lambda(b)}(X) \leqslant b$, and thus

$$\inf_{\operatorname{VaR}_{A(x)}(X) < M} \left\{ \operatorname{VaR}_{A(x)}(X) \vee x \right\} = \inf_{\operatorname{VaR}_{A(x)}(X) < M, \ x < b} \left\{ \operatorname{VaR}_{A(x)}(X) \vee x \right\} \wedge b$$

$$\leq b < M.$$

It yields that

$$\inf_{l \in \mathbb{R}_+} \Lambda \operatorname{VaR}(T_l(X)) < M$$

with corresponding optimal retention $I^* = \infty$. This contradicts with the fact that $I^* < \infty$.

Therefore, if the solution $0 < l^* < \infty$ for (11) exists, then $\theta^* > F_X(0)$ and $\Lambda(M-\epsilon) \geqslant F_X(M-\epsilon)$ for all $\epsilon \in (0,M]$. This completes the proof. \square

Although the inner problem in (11) and the optimal deductible in Theorem 1 are similar to the results for VaR-based optimal insurance problems established in Cai and Tan (2007) and Chi and Tan (2011), Theorem 2 provides a non-trivial characterization result for the existence of a positive optimal deductible, which is clearly more general than that for the VaR problem. Specifically, if $\Lambda(x) \equiv \alpha$, the second condition in Theorem 2 simplifies to $F(M) < \alpha$, implying $M < \text{VaR}_{\alpha}(X)$. This directly aligns with the result in Theorem 2.1 of Cai and Tan (2007), which is a corollary of Theorems 1 and 2.

Corollary 1. If $\Lambda(x) \equiv \alpha$ for any $x \geqslant 0$, where $\alpha \in (0,1)$, the optimal deductible $l^* > 0$ for the problem (11) exists if and only if both conditions $\theta^* > F_X(0)$ and $M < \operatorname{VaR}_{\alpha}(X)$ are satisfied. Furthermore, if l^* exists, then it holds that $l^* = d^*$.

3.2. Optimal retention with general contract

In this section, we explicitly derive the optimal solution for the ceded loss functions in the general form defined by (3). Under the expected value premium principle, we aim to solve the following problem:

$$\inf_{f \in \mathcal{F}} \Lambda \text{VaR}(X - f(X) + (1 + \theta)\mathbb{E}[f(X)]). \tag{16}$$

For $x \in \mathbb{R}_+$, denoted by

$$G(x) = d^* \wedge \operatorname{VaR}_{A(x)}(X) + (1 + \theta) \mathbb{E}\left[\min\left\{\left(X - d^*\right)_+, \left(\operatorname{VaR}_{A(x)}(X) - d^*\right)_+\right\}\right].$$
 (17)

Theorem 3. Let $\Lambda \in \mathcal{D}$, and suppose that Λ is right-continuous. The optimal f^* that solves the Λ VaR-based insurance model (16) over the class of ceded loss functions \mathcal{F} is given by

$$f^*(x) = \begin{cases} \min\left\{(x-d^*)_+, \operatorname{VaR}_{A(x^*)}(X) - d^*\right\}, & \text{if } d^* < \operatorname{VaR}_{A(x^*)}(X), \\ 0, & \text{otherwise} \end{cases}$$

(18)

in which

$$x^* = \inf\{x \in \mathbb{R}_+ : G(x) \le x\} < \infty. \tag{19}$$

Moreover, $\Lambda VaR(T_{f^*}(X)) = x^*$.

Proof. (i) Given $x \in \mathbb{R}_+$, we first solve the inner minimization of (7), that is, $\inf_{f \in F} g(x)$, where

$$g(x) := \operatorname{VaR}_{\Lambda(x)}(X) - \operatorname{VaR}_{\Lambda(x)}(f(X)) + \Pi(f(X)).$$

We can rewrite g(x) as

$$g(x) = \int_0^\infty h(S_X(y)) \,\mathrm{d}y + \int_0^\infty ((1+\theta)S_X(y) - h(S_X(y))) q(y) \,\mathrm{d}y,$$

where $h(t) = \mathbb{1}_{\{t > 1 - A(x)\}}$, and it is clear that the following function q minimizes g:

$$q(y) = \begin{cases} 0, & \text{if } (1+\theta)S_X(y) - h(S_X(y)) > 0, \\ 1, & \text{if } (1+\theta)S_X(y) - h(S_X(y)) < 0, \\ c(y), & \text{otherwise,} \end{cases}$$
 (20)

where c could be any [0,1]-valued Lebesgue-measurable function on $\{y: (1+\theta)S_X(y) - h(S_X(y)) = 0\}$. Define

$$H(y) = (1+\theta)S_X(y) - h(S_X(y)).$$

To give an explicit form of q, it suffices to study the sign of H. Noting that $S_X(y) > 1 - \Lambda(x)$ if and only if $y < \text{VaR}_{\Lambda(x)}(X)$, we have

$$H(y) = \begin{cases} (1+\theta)S_X(y) - 1, & \text{if } y < \mathrm{VaR}_{\varLambda(x)}(X), \\ (1+\theta)S_X(y), & \text{if } y \geqslant \mathrm{VaR}_{\varLambda(x)}(X). \end{cases}$$

Recall that $d^* = \operatorname{VaR}_{\theta^*}(X)$ where $\theta^* = \theta/(1+\theta)$. When $y < \operatorname{VaR}_{A(x)}(X)$, H(y) > 0 if $y \in [0, \operatorname{VaR}_{A(x)}(X) \wedge d^*)$ and $H(y) \leqslant 0$ if $y \in [\operatorname{VaR}_{A(x)}(X) \wedge d^*, \operatorname{VaR}_{A(x)}(X))$. When $y \geqslant \operatorname{VaR}_{A(x)}(X)$, $H(y) \geqslant 0$. Thus, we can set q(y) = 0 for $y \in [0, d^* \wedge \operatorname{VaR}_{A(x)}(X)) \cup [\operatorname{VaR}_{A(x)}(X), \infty)$, and q(y) = 1 for $y \in [\operatorname{VaR}_{A(x)}(X) \wedge d^*, \operatorname{VaR}_{A(x)}(X))$, which implies $f^*(y) = \min \left\{ (y - d^*)_+, (\operatorname{VaR}_{A(x)}(X) - d^*)_+ \right\}$. Therefore, we can obtain that $\inf_{f \in F} g(x) = G(x)$.

(ii) We are now ready to solve the outer optimization problem, that is,

$$\inf_{f \in \mathcal{F}} \Lambda \operatorname{VaR}(T_f(X)) = \inf_{x \in \mathbb{R}_+} \left\{ G(x) \lor x \right\}.$$

It is evident that G(x) decreases in x as $\operatorname{VaR}_{\Lambda(x)}(X)$ decreases in x. Define

$$\bar{x} = \sup\{x \in \mathbb{R}_+ : \text{VaR}_{A(x)}(X) > d^*\}, \text{ and } x^* = \inf\{x \in \mathbb{R}_+ : G(x) \le x\}.$$

Note that \bar{x} can be ∞ . Further, as $\Lambda(0) < 1$ and $\mathbb{P}(X < \infty) = 1$, we have $x^* < \infty$. Let

$$G_1(x) = d^* + (1+\theta)\mathbb{E}\left[\min\left\{\left(X - d^*\right)_+, (\text{VaR}_{A(x)}(X) - d^*)_+\right\}\right].$$

Since $\Lambda(x)$ is right-continuous in x, we have $G(x)=G_1(x)$ for $x<\overline{x}$ and $G(x)=\operatorname{VaR}_{\Lambda(x)}$ for $x\geqslant \overline{x}$. Additionally, if $x<\overline{x}$, we have $d^*<\operatorname{VaR}_{\Lambda(x)}(X)$, then

$$G(x) = d^* + (1 + \theta) \mathbb{E} \left[\min \left\{ \left(X - d^* \right)_+, \left(\text{VaR}_{A(x)}(X) - d^* \right)_+ \right\} \right]$$

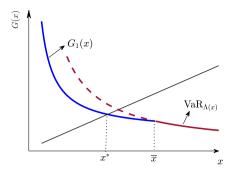
$$< d^* + (1 + \theta) \left(\text{VaR}_{A(x)}(X) - d^* \right) (1 - F(d^*)) \leqslant \text{VaR}_{A(x)}(X).$$

The second inequality is because $F(d^*) = \theta^* = \theta/(1+\theta)$. Thus, if $x^* < \overline{x}$, $\inf_{x \in \mathbb{R}_+} \{G(x) \lor x\} = x^* < \Lambda \text{VaR}(X)$, and if $x^* \geqslant \overline{x}$, $\inf_{x \in \mathbb{R}_+} \{G(x) \lor x\} = x^* = \Lambda \text{VaR}(X)$. The relationship between G(x) and x in the proof above is illustrated clearly in Fig. 1. As Λ is right-continuous, the infimum is attainable, thus $\min_{x \in \mathbb{R}_+} \{G(x) \lor x\} = x^*$ and the minimum is attained at $x = x^*$. Therefore, according to (i), we have $\inf_{f \in \mathcal{F}} \Lambda \text{VaR}(T_f(X)) = x^*$.

Remark 3. Theorem 3.2 of Chi and Tan (2011) provides the optimal ceded loss function for the case of VaR. Thus, for the first part of Theorem 3, we can directly apply the result from Chi and Tan (2011) to the inner minimization problem and solve the outer optimization problem accordingly. However, in this paper, we use a different method based on the marginal indemnity function, and therefore, we provide a separate proof.

Additionally, by comparing x^* in Theorem 3 with M in Theorem 1, and noting that $M \ge G(x)$ for any $x \in \mathbb{R}_+$, we obtain

$$\Lambda \text{VaR}\left(T_{l^*}(X)\right) \geqslant \Lambda \text{VaR}\left(T_{f^*}(X)\right)$$
.



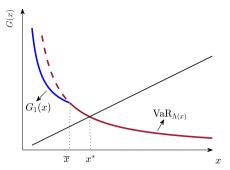


Fig. 1. An illustration of the relationship between G(x) and x.

This result is expected, as the admissible set of f includes the special form of stop-loss insurance contracts. In contrast, for a general ceded loss function, the optimal form is a limited stop-loss with an upper limit of $\operatorname{VaR}_{A(x^*)}(X) - d^*$. Finally, if $\theta = 0$, meaning that the premium corresponds to a cheap insurance, the DM always transfers all the risk to the insurer when $x < \operatorname{VaR}_{A(x^*)}(X)$.

In the following corollary, we present a special case where $\Lambda \mathrm{VaR}$ simplifies to VaR .

Corollary 2. If $\Lambda(x) \equiv \alpha$ for any $x \geqslant 0$, where $\alpha \in (0,1)$, the solution to (5) is given by

$$f^*(x) = \begin{cases} \min\left\{(x-d^*)_+ \,, \operatorname{VaR}_\alpha(X) - d^*\right\}, & \textit{if } d^* \leqslant \operatorname{VaR}_\alpha(X), \\ 0, & \textit{otherwise} \ , \end{cases}$$

and thus

$$\begin{split} \operatorname{VaR}_{\alpha}\left(T_{f^*}(X)\right) &= d^* \wedge \operatorname{VaR}_{\alpha}(X) \\ &+ (1+\theta)\mathbb{E}\left[\min\left\{\left(X - d^*\right)_+, \left(\operatorname{VaR}_{\alpha}(X) - d^*\right)_+\right\}\right]. \end{split}$$

This is consistent with the optimal indemnity function obtained by Chi and Tan (2011).

By comparing Theorem 3 and Corollary 2, we observe that the optimal ceded loss function for $\Lambda \mathrm{VaR}$ is identical to the one for VaR when $d^* \leqslant \min(\mathrm{VaR}_\alpha(X), \mathrm{VaR}_{\Lambda(x^*)}(X))$. However, $\Lambda \mathrm{VaR}$ can offer greater flexibility compared to a VaR -based contract in determining the upper limit of the ceded loss function, as x^* depends on Λ and θ .

4. Optimal results under Λvar premium principle

This section studies the case in which a Lambda-VaR risk measure is incorporated into the premium calculation. Using a Λ VaR can be attractive for the insurer for the following two reasons. First, the VaR is often used as a quantile or percentile premium Kaas et al. (e.g., Section 5 in 2008), where the maximum probability of a loss in a contract is bounded by a fixed value. Λ VaR can be seen as an extension of the quantile premium, where the quantile parameter is not fixed but depends on the loss distribution itself. This setting allows Λ VaR to flexibly adjust the premium based on the size of the loss and the risk preferences, making it more adaptable for different risk scenarios. Second, suppose that the insurer measures the risk position after selling the insurance indemnity f(X) for the deterministic premium $\Pi \geqslant 0$ by: Λ' VaR($f(X) - \Pi$), where $\Lambda' \in \mathcal{D}$. By cash subadditivity of Λ' VaR (see Han et al., 2025), a premium of the insurer that satisfies

$$\Lambda' \text{VaR}(f(X) - \Pi) \leq \Lambda' \text{VaR}(0) = 0$$

must satisfy $\Pi \geqslant \Lambda' \mathrm{VaR}(f(X))$. Different from (9) for the expected premium principle, using the $\Lambda' \mathrm{VaR}$ premium principle mitigates the total risk position of the insurer compared with no insurance when

he measures the risk by Λ' VaR. This may be preferable for risk-averse insurers in some situations. Therefore, we define the premium principle for a general $f \in \mathcal{F}$ as

$$\Pi(f(X)) = \Lambda' \text{VaR}(f(X)),$$

and aim to solve the following optimization problem

$$\inf_{f \in \mathcal{F}} \Lambda \text{VaR}(T_f(X)) = \inf_{f \in \mathcal{F}} \Lambda \text{VaR}(X - f(X) + \Lambda' \text{VaR}(f(X)))$$
 (21)

with $\Lambda, \Lambda' \in \mathcal{D}$.

It is easy to verify that for $\alpha, \beta \in (0,1)$, $f^*(x) = x \mathbb{1}_{\{\alpha \geqslant \beta\}}$ is a solution to the following optimization problem

$$\inf_{f \in \mathcal{F}} \operatorname{VaR}_{\alpha}(X - f(X)) + \operatorname{VaR}_{\beta}(f(X)). \tag{22}$$

A natural question arises: Whether a similar observation can still be made for the optimization problem (21). Note that ΛVaR is monotonic with respect to Λ in the sense that if $\Lambda_1 \leqslant \Lambda_2$ then each of the $\Lambda_1 \text{VaR}$ is smaller than the corresponding $\Lambda_2 \text{VaR}$; see Proposition 3c of Bellini and Peri (2022).

In the following proposition, we show that for the ΛVaR premium principle, a bang–bang strategy can still be optimal.

Proposition 2. For $\Lambda, \Lambda' \in \mathcal{D}$, the optimal f^* that solves the Λ VaR-based insurance model (21) over the class of ceded loss functions \mathcal{F} is given by

$$f^{*}(x) = \begin{cases} x, & \text{if } \Lambda' \text{VaR}(X) \leq \Lambda \text{VaR}(X), \\ 0, & \text{if } \Lambda' \text{VaR}(X) > \Lambda \text{VaR}(X). \end{cases}$$
 (23)

The corresponding minimum is given by

$$\Lambda \text{VaR}\left(T_{f^*}(X)\right) = \Lambda \text{VaR}(X) \wedge \Lambda' \text{VaR}(X).$$

Proof. Note that, by Lemma 1 and Proposition 1,

$$\begin{split} \inf_{f \in \mathcal{F}} \Lambda \operatorname{VaR}(T_f(X)) &= \inf_{x \in \mathbb{R}_+} \left\{ \inf_{f \in \mathcal{F}} \left\{ \operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) \right. \right. \\ &+ \left. H(f(X)) \right\} \vee x \right\} \\ &= \inf_{x \in \mathbb{R}_+} \left\{ \inf_{f \in \mathcal{F}} \left\{ \operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) \right. \right. \\ &+ \inf_{y \in \mathbb{R}_+} \left\{ \operatorname{VaR}_{A'(y)}(f(X)) \vee y \right\} \right\} \vee x \right\} \\ &= \inf_{x \in \mathbb{R}_+} \left\{ \inf_{y \in \mathbb{R}_+} \left\{ \operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) \right. \right. \\ &+ \left. \operatorname{VaR}_{A'(y)}(f(X)) \vee y \right\} \vee x \right\}. \end{split}$$

Fix $x, y \ge 0$. If $\Lambda(x) \le \Lambda'(y)$, we have $f^*(X) = 0$. This is because for any f(X) > 0 and y such that $\Lambda(x) \le \Lambda'(y)$, we have

$$\operatorname{VaR}_{\Lambda(x)}(X) - \operatorname{VaR}_{\Lambda(x)}(f(X)) + \operatorname{VaR}_{\Lambda'(y)}(f(X)) \lor y \geqslant \operatorname{VaR}_{\Lambda(x)}(X),$$

which implies that

$$\begin{split} &\inf_{x\in\mathbb{R}_+} \left\{ \inf_{y\in\mathbb{R}_+} \left\{ \operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) + \operatorname{VaR}_{A'(y)}(f(X)) \vee y \right\} \vee x \right\} \\ &\geqslant \inf_{x\in\mathbb{R}_+} \left\{ \operatorname{VaR}_{A(x)}(X) \vee x \right\} = A \operatorname{VaR}(X). \end{split}$$

 $^{^2\,}$ Here, we extend the domain of $\varLambda'VaR$ to the set $\mathcal X,$ where \varLambda is a decreasing function defined on $\mathbb R.$

On the other hand, if $\Lambda(x) > \Lambda'(y)$, we have $f^*(X) = X$. Indeed, for any $f \in \mathcal{F}$, if $\operatorname{VaR}_{\Lambda'(y)}(f(X)) \leq y$, then

$$\operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) + \operatorname{VaR}_{A'(y)}(f(X)) \vee y$$

$$\geqslant \operatorname{VaR}_{\Lambda'(y)}(X) - \operatorname{VaR}_{\Lambda'(y)}(f(X)) + y \geqslant \operatorname{VaR}_{\Lambda'(y)}(X),$$

while if $VaR_{\Lambda'(y)}(f(X)) > y$, then

$$VaR_{\Lambda(x)}(X) - VaR_{\Lambda(x)}(f(X)) + VaR_{\Lambda'(y)}(f(X)) \vee y$$

$$\geqslant \operatorname{VaR}_{\Lambda'(\gamma)}(X) - \operatorname{VaR}_{\Lambda'(\gamma)}(f(X)) + \operatorname{VaR}_{\Lambda'(\gamma)}(f(X)) = \operatorname{VaR}_{\Lambda'(\gamma)}(X).$$

Thus, we can see that the optimal contract is either full or no insurance, and its form does not depend on x and y. Therefore, if $f^*(X) = X$, we have

$$\inf_{f \in \mathcal{F}} \Lambda \mathrm{VaR}(T_f(X)) = \inf_{x \in \mathbb{R}_+} \left\{ \inf_{y \in \mathbb{R}_+} \left\{ \ \mathrm{VaR}_{\Lambda'(y)}(X) \vee y \right\} \vee x \right\} = \Lambda' \mathrm{VaR}(X).$$

Otherwise, we have

$$\inf_{f \in \mathcal{F}} \Lambda \operatorname{VaR}(T_f(X)) = \inf_{x \in \mathbb{R}_+} \left\{ \inf_{y \in \mathbb{R}_+} \left\{ \operatorname{VaR}_{\Lambda(x)}(X) \vee y \right\} \vee x \right\} = \Lambda \operatorname{VaR}(X),$$

which completes the proof.

Remark 4. Note that the problem (22) can also be interpreted as a risk-sharing problem between two agents, with VaR serving as their preference measure due to its cash additivity property. Therefore, having one agent bear all the risk can be an optimal solution. By Theorem 11 of Xia and Hu (2024), the solution to (21) also applies to the problem of risk sharing between two Λ VaR agents, even though Λ VaR is not cash additive. In fact, Theorem 11 of Xia and Hu (2024) shows that for $\Lambda_i: \mathbb{R}_+ \to [0,1]$ being decreasing for $i \in \{1,\dots,n\}$ with $n \ge 2$ and $\Lambda = \min_{1 \le i \le m} \Lambda_i$, we have

$$\inf_{(X_1,...,X_m)\in\mathbb{A}_n(X)}\sum_{i=1}^m \Lambda_i \mathrm{VaR}(X_i) = \Lambda \mathrm{VaR}(X),$$

where

$$\mathbb{A}_n(X) = \{(X_1, \dots, X_n) : X_i \in \mathcal{X}_+, \sum_{i=1}^n X_i = X, \text{ and } X_1, \dots, X_n\}$$

are comonotonic).

For a comprehensive discussion on VaR-based and Λ VaR-based risk-sharing problems, see Embrechts et al. (2018), Liu (2024) and Xia and Hu (2024).

The optimal ceded loss function in Proposition 2 is generally not unique, similar to the case of VaR. Although a bang–bang strategy can be mathematically optimal, it is often less desirable and practically applicable in real-world settings. To address this, we modify the premium in the subsequent context as follows:

$$\Pi(f(X)) = \mathbb{E}[f(X)] + \theta(\Lambda' \text{VaR}(f(X)) - \mathbb{E}[f(X)])$$

where $\Lambda' \in \mathcal{D}$ and $\theta \in [0,1]$ represents a safety loading. Similarly to (9), such a premium principle satisfies the following condition for a risk-neutral insurer:

$$\mathbb{E}[-W + f(X) - \Pi(f(X)) + C'(f(X))] = \mathbb{E}[-W],$$

with the cost of the insurer $C'(f(X)) = \theta(\Lambda' \operatorname{VaR}(f(X)) - \mathbb{E}[f(X)])$. This formulation accounts for both the expected value of the function f(X) and an additional safety margin that scales with the deviation of $\Lambda' \operatorname{VaR}(f(X))$ from $\mathbb{E}[f(X)]$. The optimization problem in (21) becomes

$$\inf_{f \in \mathcal{F}} \Lambda \text{VaR}(X - f(X) + \mathbb{E}[f(X)] + \theta(\Lambda' \text{VaR}(f(X)) - \mathbb{E}[f(X)])). \tag{24}$$

Theorem 4. Let $\Lambda, \Lambda' \in \mathcal{D}$, and suppose that Λ is right-continuous. The optimal f^* that solves the Λ VaR-based insurance model (24) over the class of ceded loss functions \mathcal{F} is given by

$$f^*(x) = x \wedge \text{VaR}_{\Lambda(x^*)}(X), \tag{25}$$

in which

$$x^* = \inf\{x \in \mathbb{R}_+ : (1 - \theta) \mathbb{E}[X \wedge \text{VaR}_{\Lambda(x)}(X)] + \theta(\Lambda' \text{VaR}(X) \wedge \text{VaR}_{\Lambda(x)}(X)) \leq x\} < \infty.$$
(26)

Moreover, $\Lambda VaR(T_{f^*}(X)) = x^*$.

Proof. By Lemma 1 and Proposition 1, we have $\inf_{X} AVaR(T_f(X))$

$$\begin{split} &=\inf_{x\in\mathbb{R}_+}\left\{\inf_{f\in\mathcal{F}}\left\{\mathrm{VaR}_{A(x)}(X)-\mathrm{VaR}_{A(x)}(f(X))+\Pi(f(X))\right\}\vee x\right\}\\ &=\inf_{x\in\mathbb{R}_+}\left\{\inf_{f\in\mathcal{F}}\left\{\mathrm{VaR}_{A(x)}(X)-\mathrm{VaR}_{A(x)}(f(X))+\mathbb{E}[f(X)]+\theta(\Lambda'\mathrm{VaR}(f(X))\right.\\ &-\mathbb{E}[f(X)])\right\}\vee x\right\}\\ &=\inf_{x\in\mathbb{R}_+}\left\{\inf_{f\in\mathcal{F}}\left\{\mathrm{VaR}_{A(x)}(X)-\mathrm{VaR}_{A(x)}(f(X))\right.\\ &+(1-\theta)\mathbb{E}[f(X)]+\theta\inf_{y\in\mathbb{R}_+}\left\{\mathrm{VaR}_{A'(y)}(f(X))\vee y\right\}\right\}\vee x\right\}\\ &=\inf_{x\in\mathbb{R}_+}\left\{\inf_{y\in\mathbb{R}_+}\inf_{f\in\mathcal{F}}\left\{\mathrm{VaR}_{A(x)}(X)-\mathrm{VaR}_{A(x)}(f(X))\right.\\ &+(1-\theta)\mathbb{E}[f(X)]+\theta\mathrm{VaR}_{A'(y)}(f(X))\vee\theta y\right\}\vee x\right\}. \end{split}$$

We first solve the inner optimization problem for a fixed $x, y \ge 0$. Note that

$$\begin{split} &\inf_{f \in \mathcal{F}} \left\{ \mathrm{VaR}_{A(x)}(X) - \mathrm{VaR}_{A(x)}(f(X)) + (1-\theta)\mathbb{E}[f(X)] \right. \\ &+ \theta \mathrm{VaR}_{A'(y)}(f(X)) \vee \theta y \right\} \\ &= \inf_{f \in \mathcal{F}, \ y < \mathrm{VaR}_{A'(y)}(f(X))} \left\{ \mathrm{VaR}_{A(x)}(X) - \mathrm{VaR}_{A(x)}(f(X)) + (1-\theta)\mathbb{E}[f(X)] \right. \\ &+ \theta \mathrm{VaR}_{A'(y)}(f(X)) \right\} \wedge \\ &\inf_{f \in \mathcal{F}, \ y \geqslant \mathrm{VaR}_{A'(y)}(f(X))} \left\{ \mathrm{VaR}_{A(x)}(X) - \mathrm{VaR}_{A(x)}(f(X)) + (1-\theta)\mathbb{E}[f(X)] \right. \\ &+ \theta y \} \,. \end{split}$$

For a fixed $x, y \ge 0$, we solve the above two problems with constraints separately:

Step 1:

$$\inf_{f \in \mathcal{F}, y < \operatorname{VaR}_{A'(y)}(f(X))} \left\{ \operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) + (1 - \theta) \mathbb{E}[f(X)] \right.$$

$$\left. + \theta \operatorname{VaR}_{A'(y)}(f(X)) \right\}$$

$$= \inf_{f \in \mathcal{F}, \lambda \in \mathbb{R}_{+}} \left\{ \operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) + (1 - \theta) \mathbb{E}[f(X)] \right.$$

$$\left. + \theta \operatorname{VaR}_{A'(y)}(f(X)) + \lambda(y - \operatorname{VaR}_{A'(y)}(f(X))) \right\},$$

$$(27)$$

where $\lambda \ge 0$ is the Lagrange multiplier.

For
$$x, y \ge 0$$
, define
$$g_1(x, y) = \int_0^\infty h_1(S_X(w)) dw$$
$$+ \int_0^\infty ((1 - \theta)S_X(w) - h_1(S_X(w))$$
$$+ (\theta - \lambda)h_2(S_X(w)))q(w) dw + \lambda y,$$

where $h_1(t) = \mathbb{1}_{\{t > 1 - A(x)\}}$ and $h_2(t) = \mathbb{1}_{\{t > 1 - A'(y)\}}$. It is clear that the following q_1 will minimize g_1 :

$$q_1(w) = \begin{cases} 0, & \text{if } (1-\theta)S_X(w) - h_1(S_X(w)) + (\theta-\lambda)h_2(S_X(w)) > 0, \\ 1, & \text{if } (1-\theta)S_X(w) - h_1(S_X(w)) + (\theta-\lambda)h_2(S_X(w)) < 0, \\ c(w), & \text{otherwise,} \end{cases}$$

where c could be any [0,1]-valued Lebesgue-measurable function on $\{w: (1-\theta)S_X(w) - h_1(S_X(w)) + (\theta-\lambda)h_2(S_X(w)) = 0\}$. Let

$$H_1(w) = (1 - \theta)S_X(w) - h_1(S_X(w)) + (\theta - \lambda)h_2(S_X(w)).$$

Note that for a fixed $x, y \ge 0$, we have two scenarios: $\Lambda(x) \le \Lambda'(y)$ and $\Lambda(x) > \Lambda'(y)$.

Case 1: $\Lambda(x) \leqslant \Lambda'(y)$

- If $S_X(w) > 1 \Lambda(x)$, that is, $w < \text{VaR}_{\Lambda(x)}(X) \le \text{VaR}_{\Lambda'(y)}(X)$, we have $H_1(w) = (1 \theta)S_X(w) 1 + (\theta \lambda) < 0$.
- If $S_X(w) \le 1 \Lambda'(y)$, that is, $w \ge \text{VaR}_{\Lambda'(y)}(X)$, we have $H_1(w) = (1 \theta)S_X(w) \ge 0$.
- If $1-\Lambda'(y) < S_X(w) \leqslant 1-\Lambda(x)$, that is, $\operatorname{VaR}_{\Lambda(x)}(X) \leqslant w < \operatorname{VaR}_{\Lambda'(y)}(X)$, we have $H_1(w) = (1-\theta)S_X(w) + (\theta-\lambda)$. Let $d_0 = \operatorname{VaR}_{\theta_0}(X)$ with $\theta_0 = 0 \lor \frac{1-\lambda}{1-\theta} \land 1$. We have $H_1(w) \leqslant 0$ if $w \geqslant d_0$ and $H_1(w) > 0$ if $w < d_0$.

By the above analysis, we can always construct a strategy in the form of $f(w) = w \wedge \operatorname{VaR}_{A(x)}(X) + (w - \tilde{d})_+ \wedge (\operatorname{VaR}_{A'(y)}(X) - \tilde{d})$, where $\tilde{d} = d_0 \vee \operatorname{VaR}_{A(x)}(X) \wedge \operatorname{VaR}_{A'(y)}(X)$, that is better than any other $f \in \mathcal{F}$. Thus, we have reduced the infinite-dimensional problem of finding a function f to minimize (27) to the one-dimensional problem of finding the optimal value of d. That is,

$$\begin{split} &\inf_{f \in \mathcal{F}, \operatorname{VaR}_{A'(y)}(f(X)) > y} \left\{ \operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) + (1 - \theta) \mathbb{E}[f(X)] \right. \\ &+ \theta \operatorname{VaR}_{A'(y)}(f(X)) \Big\} \\ &= \inf_{\substack{\operatorname{VaR}_{A(x)}(X) \le d \le \operatorname{VaR}_{A'(y)}(X) \\ d < \operatorname{VaR}_{A'(y)}(X) + \operatorname{VaR}_{A(x)}(X) - y}} \left\{ (1 - \theta) \mathbb{E}[X \wedge \operatorname{VaR}_{A(x)}(X) \\ &+ (X - d)_{+} \wedge (\operatorname{VaR}_{A'(y)}(X) - d)] \right. \\ &+ \theta (\operatorname{VaR}_{A'(y)}(X) - d + \operatorname{VaR}_{A(x)}(X)) \Big\} \,. \end{split}$$

Denote by

$$\begin{split} L_1(d) &= (1-\theta) \mathbb{E}[(X-d)_+ \wedge (\operatorname{VaR}_{A'(y)}(X) - d)] + \theta(\operatorname{VaR}_{A'(y)}(X) - d) \\ &+ \operatorname{VaR}_{A(x)}(X)). \end{split}$$

It is clear that $L_1(d)$ is continuous in d and its first-order derivative is given by $L_1'(d) = (1-\theta)F(d) - 1 < 0$. Thus, $d^* = \operatorname{VaR}_{A'(y)}(X) + (\operatorname{VaR}_{A(x)}(X) - y)_-$ minimizes $L_1(d)$. Therefore, we have

$$f^*(w) = w \wedge \text{VaR}_{A(x)}(X) + (w - d^*)_+ \wedge (y - \text{VaR}_{A(x)}(X))_+,$$

and thus

$$G_1(x, y)$$

$$\begin{split} &=\inf_{f\in\mathcal{F},\operatorname{VaR}_{A'(y)}(f(X))>y}\left\{\operatorname{VaR}_{A(x)}(X)-\operatorname{VaR}_{A(x)}(f(X))+(1-\theta)\mathbb{E}[f(X)]\right.\\ &+\theta\operatorname{VaR}_{A'(y)}(f(X))\right\}\\ &=(1-\theta)\mathbb{E}[X\wedge\operatorname{VaR}_{A(x)}(X)+(X-d^*)_+\wedge(y-\operatorname{VaR}_{A(x)}(X))_+]\\ &+\theta(\operatorname{VaR}_{A(x)}(X)+(y-\operatorname{VaR}_{A(x)}(X))_+). \end{split}$$

Case 2: $\Lambda(x) > \Lambda'(y)$

- If $S_X(w) > 1 \Lambda'(y)$, that is, $w < \operatorname{VaR}_{\Lambda'(y)}(X) \le \operatorname{VaR}_{\Lambda(x)}(X)$, we have $H_1(w) = (1 \theta)S_X(w) 1 + (\theta \lambda)$. It is obvious that $H_1(w) < 0$.
- If $S_X(w) \le 1 \Lambda(x)$, that is, $w \ge \text{VaR}_{\Lambda(x)}(X)$, we have $H_1(w) = (1 \theta)S_X(w) \ge 0$.
- If $1 \Lambda(x) < S_X(w) \le 1 \Lambda'(y)$, that is, $VaR_{\Lambda'(y)}(X) \le w < VaR_{\Lambda(x)}(X)$, we have $H_1(w) = (1 \theta)S_X(w) 1 < 0$.

Thus, $H_1(w) < 0$ for $w \in [0, \operatorname{VaR}_{A(x)}(X))$ and $H_1(w) > 0$ for $w \in [\operatorname{VaR}_{A(x)}(X), \infty)$. In this case, we always have $f^*(w) = w \wedge \operatorname{VaR}_{A(x)}(X)$, and

$$VaR_{\Lambda'(y)}(f^*(X)) = VaR_{\Lambda'(y)}(X \wedge VaR_{\Lambda(x)}(X)) = VaR_{\Lambda'(y)}(X).$$

It then follows that

$$\begin{split} &G_2(x,y) \\ &= \inf_{f \in \mathcal{F}, \operatorname{VaR}_{A'(y)}(f(X)) > y} \left\{ \operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) + (1-\theta) \mathbb{E}[f(X)] \right. \\ &\left. + \theta \operatorname{VaR}_{A'(y)}(f(X)) \right\} \\ &= \left\{ \begin{array}{ll} (1-\theta) \mathbb{E}[X \wedge \operatorname{VaR}_{A(x)}(X)] + \theta \operatorname{VaR}_{A'(y)}(X), & \text{if } \operatorname{VaR}_{A'(y)}(X) \geq y, \\ \infty, & \text{if } \operatorname{VaR}_{A'(y)}(X) < y. \end{array} \right. \end{split}$$

Step 2: Note that

$$\begin{split} &\inf_{f \in \mathcal{F}, \operatorname{VaR}_{A'(y)}(f(X)) \leq y} \left\{ \operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) + (1-\theta)\mathbb{E}[f(X)] + \theta y \right\} \\ &= \inf_{f \in \mathcal{F}} \left\{ \operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) + (1-\theta)\mathbb{E}[f(X)] + \theta y \\ &+ \lambda (\operatorname{VaR}_{A'(y)}(f(X)) - y) \right\} \\ &= \int_{0}^{\infty} h(S_{X}(w)) \, \mathrm{d}w + \int_{0}^{\infty} ((1-\theta)S_{X}(w) - h_{1}(S_{X}(w)) \\ &+ \lambda h_{2}(S_{X}(w))q(w) \, \mathrm{d}w + \theta y - \lambda y) \\ &=: g_{2}(x, y). \end{split}$$

It is clear that the following q_2 will minimize g_2 :

$$q_2(w) = \begin{cases} 0, & \text{if } (1-\theta)S_X(w) - h_1(S_X(w)) + \lambda h_2(S_X(w)) > 0, \\ 1, & \text{if } (1-\theta)S_X(w) - h_1(S_X(w)) + \lambda h_2(S_X(w)) < 0, \\ c(w), & \text{otherwise,} \end{cases}$$

where c could be any [0,1]-valued Lebesgue-measurable function on $\{w: (1-\theta)S_X(w) - h_1(S_X(w)) + \lambda h_2(S_X(w)) = 0\}$. Define

$$H_2(w) = (1 - \theta)S_X(w) - h_1(S_X(w)) + \lambda h_2(S_X(w)).$$

Case 1: $\Lambda(x) \leq \Lambda'(y)$

- If $S_X(w) > 1 \Lambda(x)$, that is, $w < \operatorname{VaR}_{\Lambda(x)}(X) \le \operatorname{VaR}_{\Lambda'(y)}(X)$, we have $H_2(w) = (1 \theta)S_X(w) 1 + \lambda$. Let $d_0 = \operatorname{VaR}_{\theta_0}(X)$ with $\theta_0 = (\lambda \theta)/(1 \theta)$. Then if $w > d_0$, we have $H_2(w) < 0$ and if $w < d_0$, we have $H_2(w) > 0$.
- If $S_X(w) \le 1 A'(y)$, that is, $y \ge \text{VaR}_{A'(y)}(X)$, we have $H_2(w) = (1 \theta)S_X(w) \ge 0$.
- If $1 \Lambda'(y) < S_X(w) \le 1 \Lambda(x)$, that is, $\operatorname{VaR}_{\Lambda(x)}(X) \le w < \operatorname{VaR}_{\Lambda'(y)}(X)$, we have $H_2(w) = (1 \theta)S_X(w) + \lambda \ge 0$.

By the above analysis, we can always construct a strategy in the form of $f(w) = (w-d)_+ \wedge (\operatorname{VaR}_{A(x)}(X) - d)$ with $d \leq \operatorname{VaR}_{A(x)}(X)$ that is better than any other $f \in \mathcal{F}$.

Case 2: $\Lambda(x) > \Lambda'(y)$

- If $S_X(w) > 1 \Lambda'(y)$, that is, $w < \operatorname{VaR}_{\Lambda'(y)}(X) \le \operatorname{VaR}_{\Lambda(x)}(X)$, we have $H_2(w) = (1 \theta)S_X(w) 1 + \lambda$. Let $d_0 = \operatorname{VaR}_{\theta_0}(X)$ with $\theta_0 = (\lambda \theta)/(1 \theta)$. Then if $w > d_0$, we have $H_2(w) < 0$ and if $w < d_0$, we have $H_2(w) > 0$.
- If $S_X(w) \le 1 \Lambda(x)$, that is, $w \ge \text{VaR}_{\Lambda(x)}(X)$, we have $H_2(w) = (1 \theta)S_X(w) \ge 0$.
- If $1 \Lambda(x) < S_X(w) \le 1 \Lambda'(y)$, that is, $\text{VaR}_{\Lambda'(y)}(X) \le w < \text{VaR}_{\Lambda(x)}(X)$, we have $H_2(w) = (1 \theta)S_X(w) 1 < 0$.

Thus, $f(w) = (w - d)_+ \wedge (\text{VaR}_{\Lambda(x)}(X) - d)$ with $d \leq \text{VaR}_{\Lambda'(y)}(X)$.

Combining the above two cases, we have

$$\inf_{f \in \mathcal{F}, \operatorname{VaR}_{A'(y)}(f(X)) \leqslant y} \left\{ \operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) \right.$$

$$+ (1 - \theta) \mathbb{E}[f(X)] + \theta y$$

$$= \inf_{\substack{d < \operatorname{VaR}_{A'(y)}(X) \wedge \operatorname{VaR}_{A(x)}(X) \\ d \geqslant \operatorname{VaR}_{A'(y)}(X) \wedge \operatorname{VaR}_{A(x)}(X) - y}} \left\{ d + (1 - \theta) \mathbb{E}[(X - d)_{+} \right.$$

$$\wedge \left(\operatorname{VaR}_{A(x)}(X) - d \right) \right] + \theta y \right\}.$$

Denote by

$$L_2(d) = (1-\theta)\mathbb{E}[(X-d)_+ \wedge (\operatorname{VaR}_{\varLambda(x)}(X) - d)] + d.$$

By the first order condition, we have $L_2'(d) = F(d) > 0$, which implies that $d^* = \max(\operatorname{VaR}_{A(x)}(X) \wedge \operatorname{VaR}_{A'(y)}(X) - y, 0)$, and thus

$$\begin{split} f^*(w) &= (w - (\mathrm{VaR}_{\varLambda(x)}(X) \wedge \mathrm{VaR}_{\varLambda'(y)}(X) - y)_+)_+ \wedge (\mathrm{VaR}_{\varLambda(x)}(X) \\ &\quad - (\mathrm{VaR}_{\varLambda(x)}(X) \wedge \mathrm{VaR}_{\varLambda'(y)}(X) - y)_+). \end{split}$$

Note that if $y>\mathrm{VaR}_{A(x)}(X)\wedge\mathrm{VaR}_{A'(y)}(X)$, we have $f^*(w)=w\wedge\mathrm{VaR}_{A(x)}(X)$, then

$$G_3(x, y) = \inf_{f \in \mathcal{F}, \operatorname{VaR}_{A'(y)}(f(X)) < y} \left\{ \operatorname{VaR}_{A(x)}(X) - \operatorname{VaR}_{A(x)}(f(X)) + (1 - \theta) \mathbb{E}[f(X)] + \theta y \right\}$$
$$= (1 - \theta) \mathbb{E}[(X \wedge \operatorname{VaR}_{A(x)})(X)] + y;$$

$$\begin{split} \text{otherwise, for } y \leqslant \text{VaR}_{A(x)}(X) \wedge \text{VaR}_{A'(y)}(X), \text{ we have} \\ G_3(x,y) &= \inf_{f \in \mathcal{F}, \text{VaR}_{A'(y)}(f(X)) < y} \left\{ \text{VaR}_{A(x)}(X) - \text{VaR}_{A(x)}(f(X)) \right. \\ &+ (1-\theta) \mathbb{E}[f(X)] + \theta y \} \\ &= (1-\theta)(\text{VaR}_{A(x)}(X) \wedge \text{VaR}_{A'(y)}(X) - y) \\ &+ \theta(\text{VaR}_{A(x)}(X) \wedge \text{VaR}_{A'(y)}(X)) \\ &+ (1-\theta) \mathbb{E}[(X-(\text{VaR}_{A(x)}(X) \wedge \text{VaR}_{A'(y)}(X) - y))_+ \wedge \\ & (\text{VaR}_{A(x)}(X) - (\text{VaR}_{A(x)}(X) \wedge \text{VaR}_{A'(y)}(X) - y))]. \end{split}$$

Now that the inner optimization problem has been solved, we can proceed to solve the outer optimization problem concerning y, based on the results from Steps 1 and 2. Specifically, we have the following expression:

$$\inf_{f \in \mathcal{F}} \Lambda \text{VaR}(T_f(X))$$

$$\begin{split} &=\inf_{x\in\mathbb{R}_+}\left\{\left\{\inf_{y\in\mathbb{R}_+,\Lambda(x)\leqslant A'(y)}\left\{G_1(x,y)\right\}\right.\right.\\ &\left.\wedge\inf_{y\leqslant \operatorname{VaR}_{A'(y)}(X),\Lambda(x)>A'(y)}\left\{G_2(x,y)\right\}\wedge\inf_{y\in\mathbb{R}_+}\left\{G_3(x,y)\right\}\right\}\vee x\right\}. \end{split}$$

It is evident that $G_1(x, y)$ increases in y, while $G_2(x, y)$ decreases in y. Thus, if there exists y such that $\Lambda(x) \leq \Lambda'(y)$, we have

$$\inf_{y\in\mathbb{R}_+,A(x)\leqslant A'(y)}G_1(x,y)=G_1(x,0)=(1-\theta)\mathbb{E}(X\wedge \mathrm{VaR}_{A(x)}(X))+\theta\mathrm{VaR}_{A(x)}(X),$$

otherwise, $\inf_{y \in \mathbb{R}_+, \Lambda(x) \leq \Lambda'(y)} G_1(x, y) = \infty$. Similarly, if there exists y such that $\Lambda(x) > \Lambda'(y)$, then

$$\begin{split} \inf_{y \leqslant \operatorname{VaR}_{A'(y)(X)}, A(x) > A(y)} G_2(x,y) &= G_2(x, \Lambda' \operatorname{VaR}(X)) \\ &= (1-\theta) \mathbb{E}(X \wedge \operatorname{VaR}_{A(x)}(X)) + \theta \Lambda' \operatorname{VaR}(X), \end{split}$$

otherwise, $\inf_{y \leq \operatorname{VaR}_{A'(y)}(X)} A_{(x)} > A_{(y)} G_2(x, y) = \infty$. Next, consider the minimization problem for G_3 . If $y > \operatorname{VaR}_{A(x)}(X) \wedge \operatorname{VaR}_{A'(y)}(X)$, then

$$\inf_{y > \operatorname{VaR}_{A(x)}(X) \wedge \operatorname{VaR}_{A'(y)}(X)} G_3(x, y) = G_3(x, \operatorname{VaR}_{A(x)}(X) \wedge A' \operatorname{VaR}(X))$$
$$= (1 - \theta) \mathbb{E}[(X \wedge \operatorname{VaR}_{A(x)}(X))]$$

$$= (1 - \theta)\mathbb{E}[(X \wedge \text{VaR}_{A(x)}(X))] + \text{VaR}_{A(x)}(X) \wedge \Lambda' \text{VaR}(X).$$

Otherwise, with a slight abuse of notation, let $l=\mathrm{VaR}_{A(x)}(X)\wedge\mathrm{VaR}_{A'(y)}(X)-y$ and

$$\bar{G}_3(l) = (1 - \theta) \left\{ l + \mathbb{E}[(X - l)_+ \wedge (\text{VaR}_{A(x)}(X)] - l) \right\}.$$

It is clear that $\bar{G}_3'(l) = F(l) > 0$, which implies that $G_3(x,y)$ decreases in y. In particular, when $y = \operatorname{VaR}_{A(x)}(X) \wedge \operatorname{VaR}_{A'(y)}(X)$, that is $y = \operatorname{VaR}_{A(x)}(X) \wedge A'\operatorname{VaR}(X)$, we have

$$\inf_{y \leq \operatorname{VaR}_{A(x)}(X) \wedge \operatorname{VaR}_{A'(y)}(X)} \left\{ G_3(x, y) \right\} = G_3(x, \operatorname{VaR}_{A(x)}(X) \wedge \Lambda' \operatorname{VaR}(X))$$

$$= (1 - \theta) \mathbb{E}[X \wedge \operatorname{VaR}_{A(x)}(X)]$$

$$+ \theta(\operatorname{VaR}_{A(x)}(X) \wedge \Lambda' \operatorname{VaR}(X)).$$

To summarize, we have the following result

$$\inf_{f \in \mathcal{F}} \Lambda \operatorname{VaR}(T_f(X)) = \inf_{x \in \mathbb{R}_+} \left\{ \left\{ (1 - \theta) \mathbb{E}[X \wedge \operatorname{VaR}_{\Lambda(x)}(X)] + \theta(\Lambda' \operatorname{VaR}(X) \wedge \operatorname{VaR}_{\Lambda(x)}(X)) \right\} \vee x \right\}.$$
(29)

Let

$$x^* = \inf\{x \in \mathbb{R}_+ : (1 - \theta)\mathbb{E}[X \wedge \text{VaR}_{\Lambda(x)}(X)] + \theta(\Lambda' \text{VaR}(X) \wedge \text{VaR}_{\Lambda(x)}(X)) \\ \leq x\}.$$

Similarly to the proof of Theorem 3, the function

$$x \mapsto (1 - \theta) \mathbb{E}[X \wedge \text{VaR}_{\Lambda(x)}(X)] + \theta(\Lambda' \text{VaR}(X) \wedge \text{VaR}_{\Lambda(x)}(X))$$

is decreasing in x. As $\Lambda(0) < 1$ and $\mathbb{P}(X < \infty) = 1$, we have $x^* < \infty$. Because Λ is right-continuous, the infimum in (29) is attainable and thus

$$x^* = \min_{x \in \mathbb{R}_+} \left\{ \left\{ (1 - \theta) \mathbb{E}[X \wedge \text{VaR}_{A(x)}(X)] + \theta(\Lambda' \text{VaR}(X) \wedge \text{VaR}_{A(x)}(X)) \right\} \vee x \right\},$$

which is obtained at $x=x^*$. Therefore, we have $\inf_{f\in\mathcal{F}} \Lambda \mathrm{VaR}(T_f(X))=x^*$. This completes the proof. $\ \ \, \Box$

Eq. (25) shows that the ceded loss function takes the form of a limited loss transform. Under this strategy, the DM retains any loss exceeding the threshold $\text{VaR}_{\Lambda(x^*)}(X)$. While this threshold is determined by the function Λ , the value of x^* in (26) is determined based on $\Lambda'\text{VaR}$. Notably, as $\Lambda'\text{VaR}(X)$ increases, x^* also increases, which reduces the threshold. Consequently, a higher premium results in the DM retaining a larger portion of the claim. Further numerical analysis can be found in Section 6.

The following corollary directly follows from Theorem 4.

Corollary 3. If $\Lambda(x) \equiv \alpha$ and $\Lambda'(y) \equiv \beta$ for any $x, y \geqslant 0$, where $\alpha, \beta \in (0, 1)$, the solution to (24) is given by

$$f^*(w) = w \wedge \operatorname{VaR}_{\alpha}(X),$$

and thus

$$\inf_{f \in \mathcal{F}} \operatorname{VaR}_{\alpha}(T_f(X)) = (1 - \theta) \mathbb{E}[X \wedge \operatorname{VaR}_{\alpha}(X)] + \theta(\operatorname{VaR}_{\beta}(X) \wedge \operatorname{VaR}_{\alpha}(X)).$$

5. Optimal insurance under model uncertainty

In the above sections, we assumed that the distribution of the underlying loss random variable is either given or known. However, this assumption is sometimes unrealistic because the actual loss distribution is generally unavailable in most practical scenarios. This recognition prompts us to explore model uncertainty in the determination of optimal insurance policies. In the following context, we investigate an optimal insurance problem incorporating a likelihood ratio uncertainty or assuming that the loss is partially known in the sense that only the first two moments of the loss are available. The closed-form solutions are obtained for both uncertainty scenarios. The premium is calculated using (8), which is based on the expected value premium principle.

5.1. Likelihood ratio uncertainty

Suppose the DM's reference probability measure \mathbb{P} is fixed, and Λ VaR is evaluated under \mathbb{P} . However, the true probability measure and the corresponding distribution of X might not be known. To account for this uncertainty, we use the likelihood ratio to describe the DM's attitude towards uncertainty in probability measures, as introduced in Liu et al. (2022) and Landriault et al. (2024).

We define $\mathcal P$ as the collection of probability measures that are absolutely continuous with respect to $\mathbb P$. The uncertainty set is defined as

$$\mathcal{P}_{\beta} = \left\{ \mathbb{Q} \in \mathcal{P} : \frac{d\mathbb{Q}}{d\mathbb{P}} \leq \frac{1}{\beta} \right\},\,$$

for a given $\beta \in (0,1]$ representing the degree of uncertainty faced by the DM. A larger value of β corresponds to a lower degree of uncertainty, with $\beta=1$ indicating that the DM encounters no uncertainty.

The DM aims to solve the following robust optimization problem

$$\min_{f \in \mathcal{F}} \sup_{\mathbb{Q} \in \mathcal{P}_{\beta}} \Lambda \operatorname{VaR}^{\mathbb{Q}}(T_f), \tag{30}$$

where $\Lambda VaR^{\mathbb{Q}}$ is ΛVaR evaluated under the probability measure \mathbb{Q} instead of \mathbb{P} .

Lemma 2 (Proposition 4 of Xia & Hu, 2024). Let $\Lambda: \mathbb{R} \to [0,1]$ be decreasing. For $\beta \in (0,1]$, define $\Lambda_{\beta} = \beta \Lambda + 1 - \beta$. Then

$$\sup_{\mathbb{Q}\in\mathcal{P}_{\beta}}\Lambda\mathrm{VaR}^{\mathbb{Q}}(X)=\Lambda_{\beta}\mathrm{VaR}^{\mathbb{Q}}(X),\quad X\in\mathcal{X}.$$

The following theorem is readily derived by combining Lemma 2 and Theorem 3. Specifically, it can be obtained by substituting Λ with Λ_{β} throughout Theorem 3. Recall that $d^* = \text{VaR}_{\theta^*}(X)$, where $\theta^* = \theta/(1+\theta)$.

Theorem 5. Let $\Lambda \in \mathcal{D}$ and suppose that Λ is right-continuous. For $\beta \in (0, 1]$, define $\Lambda_{\beta} = \beta \Lambda + 1 - \beta$. The optimal f^* that solves the Λ VaR-based insurance model (30) over the class of ceded loss functions \mathcal{F} is given by

$$f^*(x) = \begin{cases} \min\left\{ (x - d^*)_+, \operatorname{VaR}_{\Lambda_{\beta}(x^*)}(X) - d^* \right\}, & \text{if } d^* \leqslant \operatorname{VaR}_{\Lambda_{\beta}(x^*)}(X), \\ 0, & \text{otherwise} \end{cases}$$

in which $x^* = \inf\{x \in \mathbb{R}_+ : G(x) \leq x\}$ with

$$\begin{split} G(x) &= d^* \wedge \mathrm{VaR}_{A_{\tilde{\beta}}(x)}(X) \\ &+ (1+\theta) \mathbb{E} \left[\min \left\{ \left(X - d^* \right)_+, \left(\mathrm{VaR}_{A_{\tilde{\beta}}(x)}(X) - d^* \right)_+ \right\} \right]. \end{split}$$

Moreover, $\Lambda VaR^{\mathbb{Q}^*}(T_{f^*}(X)) = x^*$.

Note that $\Lambda_{\beta} = \Lambda + (1 - \Lambda)(1 - \beta) \geqslant \Lambda$. In particular, if $\beta = 1$, then the robust optimization problem (30) reduces to (5). As β decreases, Λ_{β} clearly increases, indicating that the DM applies a higher overall confidence level function when determining the optimal insurer strategy. Furthermore, from the equation for solving x^* , it shows that x^* increases as β decreases, due to the increases in Λ_{β} . Since Λ_{β} is a decreasing function of x, this results in a trade-off between risk and uncertainty for the DM when using Λ VaR as the measure. We provide a numerical example of $\Lambda(x^*)$ related to β in Section 6.

Remark 5. If we assume that the contract follows a stop-loss form as in Section 3.1, a solution similar to that in Theorem 1 can also be obtained. We only need to replace all instances of Λ with Λ_{β} in Theorem 1, so we do not repeat the details here.

We immediately have the following corollary, which is also shown in Example 3.1 of Landriault et al. (2024).

Corollary 4. If $\Lambda(x) \equiv \alpha$ for any $x \in \mathbb{R}_+$, where $\alpha \in (0,1)$, the optimization problem (30) can be solved by

$$f^*(x) = \begin{cases} \min\{(x - d^*)_+, \operatorname{VaR}_{\beta\alpha + 1 - \beta}(X) - d^*\}, & \text{if } d^* \leqslant \operatorname{VaR}_{\beta\alpha + 1 - \beta}(X) \\ 0, & \text{otherwise,} \end{cases}$$

and

$$\min_{f \in \mathcal{F}} \sup_{\mathbb{Q} \in \mathcal{P}_{\mathcal{B}}} \mathrm{VaR}_{\alpha}^{\mathbb{Q}}(T_f)$$

$$\begin{split} &= d^* \wedge \mathrm{VaR}_{\beta\alpha+1-\beta}(X) + (1+\theta) \mathbb{E}\left[\min\left\{\left(X-d^*\right)_+, \right. \\ &\left.\left(\mathrm{VaR}_{\beta\alpha+1-\beta}(X)-d^*\right)_+\right\}\right]. \end{split}$$

Note that $\beta\alpha+1-\beta=\alpha+(1-\alpha)(1-\beta)\geqslant\alpha$. Corollary 4 shows that the deductible in the worst-case scenario is greater than that in Corollary 2 where model uncertainty is not involved, which is a reasonable outcome given the increased risk exposure. This behavior reflects a risk-averse strategy: when faced with uncertainty about future outcomes, the DM opts for more insurance coverage to safeguard against potential adverse events. Additionally, as β decreases, $\beta\alpha+1-\beta$ increases, leading to a larger $\mathrm{VaR}_{\beta\alpha+1-\beta}(X)$. This finding suggests that as ambiguity in the probability measure increases (indicated by a higher $1/\beta$), the DM is more likely to choose higher insurance coverage.

5.2. Mean-variance uncertainty

In this subsection, we assume that the loss is partially known in the sense that only the first two moments of X are known and finite. Let \mathcal{M}_+ be the set of CDFs of random variables in \mathcal{X}_+ . Given a pair of non-negative mean and standard deviation (μ,σ) of X, we define the uncertainty set

$$S(\mu, \sigma) = \left\{ F \in \mathcal{M}_{+} : \int_{0}^{\infty} x \, dF(x) = \mu, \int_{0}^{\infty} x^{2} \, dF(x) = \mu^{2} + \sigma^{2} \right\}.$$
(31)

The $S(\mu, \sigma)$ specifies a set of CDFs with the same mean μ and variance σ^2 . The distributionally robust insurance problem is defined by

$$\inf_{f \in \mathcal{F}} \sup_{F \in \mathcal{S}(\mu, \sigma)} \Lambda \text{VaR}(T_f(X)), \tag{32}$$

where T_f is given by (4).

For a set of distributions \mathcal{M} , we denote $G_{\mathcal{M}}^-(x) = \inf_{G \in \mathcal{M}} G(x)$, for $x \in \mathbb{R}_+$. In the following context, we abuse the notation by treating ΛVaR also as a mapping from a set \mathcal{M} of distributions to \mathbb{R} ; that is, we write $\Lambda \text{VaR}(G) = \Lambda \text{VaR}(Y)$ when Y follows the distribution G.

The following lemma plays a key role in solving the problem (32).

Lemma 3 (Theorem 2 of Han & Liu, 2024). Let \mathcal{M} be a set of distributions and $\Lambda \in \mathcal{D}$. We have $\sup_{G \in \mathcal{M}} \Lambda \operatorname{VaR}(G) = \Lambda \operatorname{VaR}(G_{\mathcal{M}}^{-})$.

Proposition 3. For a set of distributions \mathcal{M} that includes the candidate distribution of X and $\Lambda \in \mathcal{D}$, we have

$$\inf_{f \in \mathcal{F}} \sup_{F \in \mathcal{M}} \Lambda \mathrm{VaR}(T_f(X)) = \inf_{x \in \mathbb{R}_+} \left\{ \inf_{f \in \mathcal{F}} \sup_{F \in \mathcal{M}} \mathrm{VaR}_{\Lambda(x)}(T_f(X)) \vee x \right\}.$$

Proof. By Lemma 3, for $d \ge 0$, we have

$$\sup_{F\in\mathcal{M}} \Lambda \mathrm{VaR}(T_f(X)) = \sup_{G\in\mathcal{M}_f} \Lambda \mathrm{VaR}(G) = \Lambda \mathrm{VaR}(G_{\mathcal{M}_f}^-),$$

where

$$\mathcal{M}_f = \{G : [0, \infty) \to [0, 1] : G \text{ is the distribution}$$

function of $T_f, F \in \mathcal{M}\}.$

Because $\sup_{F\in\mathcal{M}}\operatorname{VaR}_{A(x)}(T_f(X))=\sup_{G\in\mathcal{M}_f}\operatorname{VaR}_{\alpha}(G)=\operatorname{VaR}_{\alpha}(G_{\mathcal{M}_f}^-)$ (see Corollary 1 of Han & Liu, 2024), by Lemma 1, we have

$$\begin{split} \inf_{f \in \mathcal{F}} \sup_{F \in \mathcal{M}} \Lambda \mathrm{VaR}(T_f(X)) &= \inf_{f \in \mathcal{F}} \Lambda \mathrm{VaR}(G_{\mathcal{M}_f}^-) \\ &= \inf_{f \in \mathcal{F}} \inf_{x \in \mathbb{R}_+} \left\{ \mathrm{VaR}_{\Lambda(x)}(G_{\mathcal{M}_f}^-) \vee x \right\} \\ &= \inf_{x \in \mathbb{R}_+} \inf_{f \in \mathcal{F}} \left\{ \sup_{F \in \mathcal{M}} \mathrm{VaR}_{\Lambda(x)}(T_f(X)) \vee x \right\} \\ &= \inf_{x \in \mathbb{R}_+} \left\{ \inf_{f \in \mathcal{F}} \sup_{F \in \mathcal{M}} \mathrm{VaR}_{\Lambda(x)}(T_f(X)) \vee x \right\}, \end{split}$$

which completes the proof. \Box

In fact, Proposition 3 shows that we can switch the order of the supremum and infimum in the problem (32) as follows:

$$\begin{split} &\inf_{f \in \mathcal{F}} \sup_{F \in \mathcal{M}} \inf_{x \in \mathbb{R}_+} \left\{ \mathrm{VaR}_{A(x)}(T_f(X)) \vee x \right\} \\ &= \inf_{x \in \mathbb{R}_+} \left\{ \inf_{f \in \mathcal{F}} \sup_{F \in \mathcal{M}} \mathrm{VaR}_{A(x)}(T_f(X)) \vee x \right\}. \end{split}$$

It is important to note that the conclusion in Proposition 3 is independent of specific forms of \mathcal{M} ; it does not need to be the uncertainty set defined by $S_{\mu,\sigma}$ in (31). It can also be applied to other types of uncertainty sets, such as the Wasserstein uncertainty set or risk aggregation sets; see Section 4 of Han and Liu (2024). Also, for $\alpha \in (0,1)$, the optimization problem

$$\inf_{f \in \mathcal{F}} \sup_{F \in \mathcal{M}} \operatorname{VaR}_{\alpha}(T_f(X)) \tag{33}$$

represents the robust insurance problem with VaR. This problem has been studied in the literature; see e.g., Liu and Mao (2022) and Cai et al. (2024).

As in Section 3.1, we restrict the set of admissible insurance indemnities to be the set of stop-loss contracts, that is, we assume $f(x) = (x - l)_+$ for some $l \ge 0$. Thus, the insurer's optimization problem (32) becomes

$$\min_{l\geqslant 0}\sup_{F\in\mathcal{S}_{\mu,\sigma}}\Lambda\mathrm{VaR}(T_l(X)), \tag{34}$$

where T_l is defined in (10).

For the special case where $\Lambda(x) \equiv \alpha$ for any $x \in \mathbb{R}_+$, the optimization problem (34) reduces to the robust insurance problem with VaR, which has been addressed by Theorem 2 of Liu and Mao (2022).

Lemma 4 (Theorem 2 of Liu & Mao, 2022). Let $\Lambda(x) \equiv \alpha \in (0,1)$. If $\theta \leqslant \sigma^2/\mu^2$, then an optimal deductible of the optimization problem (34) is $l^* = \infty \mathbb{1}_{\{\alpha < \theta^*\}}$, and the optimal value is

$$\inf_{l\geqslant 0}\sup_{F\in S(\mu,\sigma)}\mathrm{VaR}_{\alpha}(T_l(X)) = \begin{cases} (1+\theta)\mu, & \text{if }\alpha\geqslant \theta^*,\\ \frac{\mu}{1-\sigma}, & \text{if }\alpha<\theta^*. \end{cases}$$

If $\theta > \sigma^2/\mu^2$, then the optimal deductible of the optimization problem (34) is

$$I^* = \left(\mu - \sigma \frac{1-\theta}{2\sqrt{\theta}}\right) \mathbbm{1}_{\{\alpha \geqslant \theta^*\}} + \infty \mathbbm{1}_{\{\alpha < \theta^*\}},$$

and the optimal value is

$$\inf_{l\geqslant 0}\sup_{F\in S(\mu,\sigma)}\mathrm{VaR}_{\alpha}(T_l(X))=\sup_{F\in S(\mu,\sigma)}\mathrm{VaR}_{\alpha}(X)=\begin{cases} \mu+\sigma\sqrt{\theta}, & \text{if }\alpha\geqslant \theta^*,\\ \mu+\sigma\sqrt{\frac{\alpha}{1-\alpha}}, & \text{if }\alpha<\theta^*. \end{cases}$$

Combining Proposition 3 and Lemma 4, we have the following theorem.

Theorem 6. Let $\Lambda \in \mathcal{D}$ and suppose that Λ is right-continuous. If $\theta \leq \sigma^2/\mu^2$, the solution $l^* \in [0, \infty]$ that solves the robust Λ VaR-based insurance model (34) is given by

$$l^* = \begin{cases} 0, & \text{if } x_1^* = (1+\theta)\mu, \\ \infty, & \text{if } x_1^* < (1+\theta)\mu, \end{cases}$$
 (35)

and the corresponding minimum is given by $\Lambda VaR\left(T_{l^*}(X)\right)=x_1^*$ with x_1^* defined as

$$x_1^* = \inf \left\{ x \in \mathbb{R}_+ : (1+\theta)\mu \mathbb{1}_{\{\Lambda(x) \geqslant \theta^*\}} + \frac{\mu}{1-\Lambda(x)} \mathbb{1}_{\{\Lambda(x) < \theta^*\}} \leqslant x \right\}.$$

If $\theta > \sigma^2/\mu^2$, the solution $l^* \in [0, \infty]$ that solves the robust Λ VaR-based insurance model (34) is given by

$$l^* = \begin{cases} \mu - \sigma \frac{1-\theta}{2\sqrt{\theta}}, & \text{if } x_2^* = \mu + \sigma\sqrt{\theta}, \\ \infty, & \text{if } x_2^* < \mu + \sigma\sqrt{\theta}, \end{cases}$$
 (36)

and the corresponding minimum is given by $\Lambda VaR\left(T_{l^*}(X)\right) = x_2^*$ with x_2^* defined as

$$\begin{split} x_2^* &= \inf \left\{ x \in \mathbb{R}_+ \, : \, \left(\mu + \sigma \sqrt{\theta} \right) \mathbb{1}_{\{\Lambda(x) \leq \theta^*\}} \right. \\ &+ \left(\mu + \sigma \sqrt{\frac{\Lambda(x)}{1 - \Lambda(x)}} \right) \mathbb{1}_{\{\Lambda(x) > \theta^*\}} \leqslant x \right\}. \end{split}$$

Proof. Recall that $\theta^* = \theta/(1+\theta)$. For $x \in \mathbb{R}_+$, $(1-\Lambda(x))(1+\theta) \le 1$ is equivalent to $\Lambda(x) \le \theta^*$. By Proposition 3 and Lemma 4, if $\theta \le \sigma^2/\mu^2$,

we have

$$\begin{split} \inf_{l \geqslant 0} \sup_{F \in S(\mu,\sigma)} \Lambda \mathrm{VaR}(T_l(X)) &= \inf_{x \in \mathbb{R}_+} \left\{ \inf_{l \geqslant 0} \sup_{F \in S(\mu,\sigma)} \mathrm{VaR}_{\Lambda(x)}(T_l(X)) \vee x \right\} \\ &= \inf_{x \in \mathbb{R}_+} \left\{ \left((1+\theta) \mu \mathbb{1}_{\{\Lambda(x) \geqslant \theta^*\}} \right. \\ &\left. + \frac{\mu}{1 - \Lambda(x)} \mathbb{1}_{\{\Lambda(x) < \theta^*\}} \right) \vee x \right\}. \end{split}$$

Recall that

$$x_1^* = \inf \left\{ x \in \mathbb{R}_+ \, : \, (1+\theta) \mu \mathbb{1}_{\{\Lambda(x) \ge \theta^*\}} + \frac{\mu}{1 - \Lambda(x)} \mathbb{1}_{\{\Lambda(x) < \theta^*\}} \le x \right\}.$$

It is clear that the function

$$x \mapsto (1+\theta)\mu \mathbb{1}_{\{\Lambda(x) \geqslant \theta^*\}} + \frac{\mu}{1-\Lambda(x)} \mathbb{1}_{\{\Lambda(x) < \theta^*\}}$$

is decreasing. Therefore, we have $x_1^* \leq (1+\theta)\mu$ and

$$\inf_{l\geqslant 0}\sup_{F\in\mathcal{S}(\mu,\sigma)}\Lambda\mathrm{VaR}(T_l(X))=x_1^*.$$

Because Λ is right-continuous, we have

$$x_1^* = \min \left\{ x \in \mathbb{R}_+ \, : \, (1+\theta)\mu \mathbb{1}_{\{\Lambda(x) \geqslant \theta^*\}} + \frac{\mu}{1-\Lambda(x)} \mathbb{1}_{\{\Lambda(x) < \theta^*\}} \leqslant x \right\}.$$

If $x_1^* = (1+\theta)\mu$, then $\Lambda(x_1^*) \ge \theta^*$. By Lemma 4, we have $l^* = 0$. If $x_1^* < (1+\theta)\mu$, then $\Lambda(x_1^*) < \theta^*$. Lemma 4 yields that $l^* = \infty$. Therefore, the optimal deductible l^* for problem (34) is given by (35). On the other hand, if $\theta > \sigma^2/\mu^2$, we have

$$\inf_{l\geqslant 0}\sup_{F\in S(\mu,\sigma)}\Lambda \mathrm{VaR}(T_l(X))$$

$$=\inf_{x\in\mathbb{R}_+}\left\{\left(\left(\mu+\sigma\sqrt{\theta}\right)\mathbb{1}_{\{\Lambda(x)\geqslant\theta^*\}}+\left(\mu+\sigma\sqrt{\frac{\Lambda(x)}{1-\Lambda(x)}}\right)\mathbb{1}_{\{\Lambda(x)<\theta^*\}}\right)\vee x\right\}.$$

Recall that

$$x_2^* = \inf \left\{ x \in \mathbb{R}_+ : \left(\mu + \sigma \sqrt{\theta} \right) \mathbb{1}_{\{\Lambda(x) \le \theta^*\}} + \left(\mu + \sigma \sqrt{\frac{\Lambda(x)}{1 - \Lambda(x)}} \right) \mathbb{1}_{\{\Lambda(x) > \theta^*\}} \right\}$$

$$\leq x \}.$$

Similarly to the first case, we have $x_2^* \le \mu + \sigma \sqrt{\theta}$ and

$$\inf_{l\geqslant 0}\sup_{F\in S(\mu,\sigma)}\Lambda\mathrm{VaR}(T_l(X))=x_2^*.$$

By the right-continuity of Λ and Lemma 4, the optimal deductible I^* is given by (36). \square

By comparing the strategies in Lemma 4 and Theorem 6, we observe that the forms of the optimal solutions are similar. However, for a VaR agent, the optimal strategy depends on the values of α and θ^* . In contrast, for a Λ VaR agent, determining the optimal strategy involves finding the optimal confidence level by solving for x_i^* (i=1,2). Some observations that hold for a VaR agent can also be applied to a Λ VaR agent. Specifically, when θ^* is higher than a certain confidence level (α for the VaR agent and $\Lambda(x_i^*)$ for the Λ VaR agent), the DM always chooses to buy no insurance. Conversely, if θ^* is smaller than the confidence level, the DM either chooses to transfer all risk to the insurer or to purchase some insurance. Notably, in (36), when $x_2^* = \mu + \sigma \sqrt{\theta}$, the optimal deductible l^* increases with θ . This suggests that the insurer opts to retain more risk as the insurance premium rises. Additionally, if $\theta > 1$, the deductible is larger than the mean of the claim due to the high cost of the premium.

6. Numerical examples

In this section, we numerically investigate how the solutions to the optimal insurance problem with ΛVaR interact with changes in the underlying model parameters.

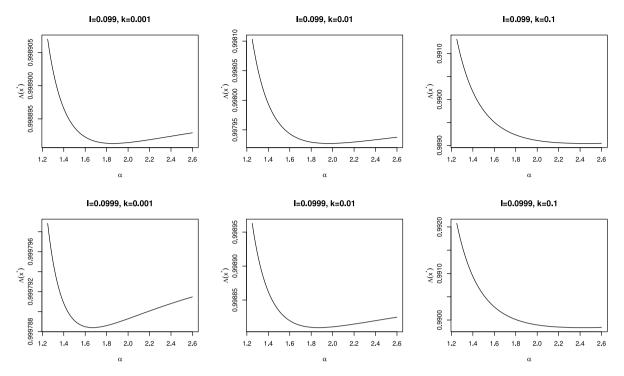


Fig. 2. $\Lambda(x^*)$ for a range of values of α in problem (5) with the expected premium principle.

Assume that the random loss X follows a Pareto distribution, which is commonly used in modeling large losses caused by catastrophic events. The Pareto distribution function with parameter $\alpha>0$ and m>0 is given by

$$\mathbb{P}(X \leqslant x) = 1 - \left(\frac{m}{x}\right)^{\alpha}, \quad x \geqslant m.$$

Here $\alpha>0$ is the tail parameter and m is the scale parameter. The Pareto distribution with a smaller α has a heavier tail. For $\alpha>1$, $\mathbb{E}[X]=\alpha m/(\alpha-1)$, and for $\alpha\leqslant 1$, $\mathbb{E}[X]=\infty$. Let

$$\Lambda(x) = \ell e^{-kx} + 0.9, \quad x \in \mathbb{R}_+,$$

where $k \ge 0$ and $\ell \in [0,0.1)$. Thus, a larger loss is tolerated with a probability level ranging from 0.9 to $0.9 + \ell$, decreasing as the loss increases. An agent with a larger ℓ is more concerned with the tail part of the risk and an agent with a larger k is more tolerant of large losses.

We first consider problem (5) with the expected premium principle and general contracts. By Theorem 3, the optimal ceded loss function is $f^*(x) = \min\{(x-d^*)_+, \operatorname{VaR}_{A(x^*)}(X) - d^*\}$, where d^* depends on θ via $d^* = \operatorname{VaR}_{\theta/(1+\theta)}(X)$ and x^* , as in (19), can be calculated numerically. Fixing $\theta = 0.25$, we compute $A(x^*)$ for different Pareto risks in cases where $k \in \{0.001, 0.01, 0.1\}$ and $\ell \in \{0.099, 0.0999\}$. Assume that the expectations of these Pareto risks are equal to 1; that is, for a fixed $\alpha > 1$, m is taken to be $1 - 1/\alpha$.

We plot $\Lambda(x^*)$ against $\alpha \in (1.25, 2.6)$ in Fig. 2. When k and ℓ are fixed, as α increases, $\Lambda(x^*)$ decreases first and it may become increasing when α reaches a certain level. In other words, when α is small, for more heavy-tailed risks, more risk needs to be transferred under the optimal strategy. This may seem counterintuitive as the DM is more tolerant of large losses. To understand this intriguing phenomenon, it is important to note that as the Pareto risks have the same expectation, one risk cannot stochastically dominate another, even if it has a heavier tail. Although heavy-tailed risks are more likely to take extremely large values, it is less probable for them to take moderately large values. Therefore, for DMs who care less about extremely large values of risks, less heavy-tailed Pareto risks can be more risky, leading to less risk

being transferred. The observation may flip when α is large, i.e., less risk is transferred for more heavy-tailed risks. This is because in these cases the DM is more concerned about the extremely large values of the risks with larger α and thus more heavy-tailed risks are indeed more dangerous. When α is fixed, DMs with larger ℓ and smaller k are more careful about extremely large values of risks, leading to more risk being transferred.

In Fig. 3, fixing $\alpha=1.5$, m=1/3, $\ell=0.09$, and k=0.1, we plot $\Lambda(x^*)$ and $\Lambda(x^*)-\theta^*$ for a range of values of θ , where $\theta^*=\theta/(1+\theta)$. From the left plot of Fig. 3, we find that less loss from the tail part will be transferred as θ gets larger. Moreover, less loss will be transferred at an overall level as well, confirmed by the right plot of Fig. 3.

Next, we continue to study problem (30) with the expected premium principle, under likelihood ratio uncertainty. For $\beta \in (0,1]$, let $\Lambda_{\beta} = \beta \Lambda + 1 - \beta$. Here, β indicates the level of uncertainty; the smaller β is, the more model uncertainty we have. By Theorem 5, the optimal ceded loss function is $f^*(x) = \min\{(x-d^*)_+, \operatorname{VaR}_{\Lambda_{\beta}(x^*)}(X) - d^*\}$. Letting $\alpha = 1.5, \ m = 1/3$, and $\theta = 0.25$, we display $\Lambda_{\beta}(x^*)$ for a range of values of β in Fig. 4. It is found that as the model uncertainty increases, more loss will be transferred. This observation is consistent with the case when VaR is used as the risk measure; see Corollary 4. As k decreases (or ℓ increases), $\Lambda_{\beta}(x^*)$ becomes larger and more risk is transferred. This is consistent with the case without model uncertainty. Moreover, when model uncertainty is large (i.e., β is small), the DM risk attitude becomes less important to their optimal decision.

Finally, we consider the case when Λ' VaR is incorporated in the premium principle (problem (24)). By Theorem 4, the optimal ceded loss function is $f^*(x) = x \wedge \text{VaR}_{\Lambda(x^*)}(X)$. We assume $\Lambda' = \Lambda$, and plot $\Lambda(x^*)$ as a function of $\alpha \in (1.05, 2.6)$ by fixing $\theta = 0.25$. We choose to plot $\Lambda(x^*)$ on a different range of α from the range used in problem (5) mainly for better illustration. The observations are similar to the case of expected value premium as shown in Fig. 5.

7. Conclusion

This paper studies the objective of minimizing the ΛVaR of terminal losses for a DM who is considering to purchase insurance. The

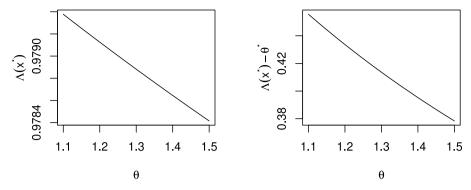


Fig. 3. $\Lambda(x^*)$ and $\Lambda(x^*) - \theta^*$ for a range of values of θ in problem (5) with the expected premium principle.

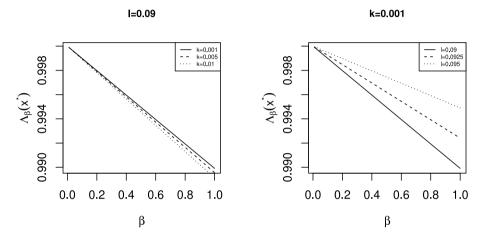


Fig. 4. $\Lambda_{\beta}(x^*)$ for a range of values of β in problem (30) under likelihood ratio uncertainty.

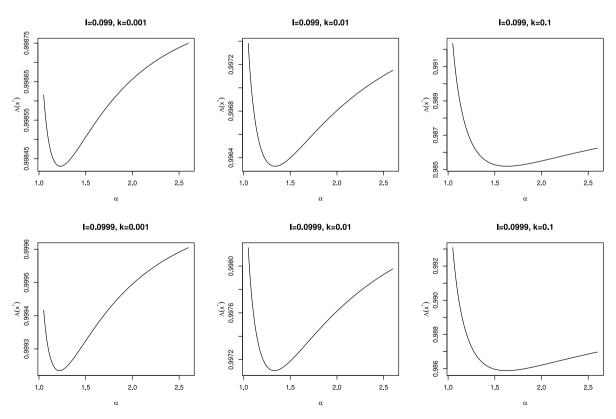


Fig. 5. $\Lambda(x^*)$ for a range of values of α in problem (24).

insurance premium is given by an expected value premium principle, or based on another Λ' VaR risk measure. Generally, a limited stoploss indemnity or full/no insurance is shown to be optimal. If the DM only considers stop-loss indemnities, then the optimal deductible is shown in closed form when the premium is given by an expected value premium principle. Finally, this paper also considers the impact of model uncertainty.

One of the core techniques we use to derive the results is that we represent ΛVaR by VaR (see Lemma 1 and Proposition 1). We convey the message that optimization problems based on ΛVaR (either maximization or minimization problems) can usually be safely transferred to equivalent VaR problems. As a result, the optimal solutions with ΛVaR and VaR seem to share similar forms. However, depending on specific practical needs, the former solution sometimes has several advantages over the latter, including that (i) the solution with ΛVaR is mathematically more general, than that with VaR as its special case; (ii) the solution with ΛVaR allows the DM to adjust her risk/confidence level along with the model parameters and insurance prices. In light of these, optimal solutions (including their existence) for ΛVaR are also non-trivial to get (see in particular Theorems 2 and 4).

While we believe that this study includes a wide class of cases, we conclude this paper with two suggestions for further research. First, this study assumes homogeneous beliefs, meaning that the premium principle and the objective function are based on the same probability measure. It would be interesting to relax this assumption, and let the premium principle be based on an alternative probability measure, just like in the risk-sharing problem of Liu et al. (2024). Second, we assume in this paper that there is only one insurer. For further research, we propose to study the impact of having multiple insurers in the market, each with a different premium principle.

CRediT authorship contribution statement

Tim J. Boonen: Writing – review & editing, Writing – original draft, Methodology. **Yuyu Chen:** Writing – review & editing, Writing – original draft, Methodology. **Xia Han:** Writing – review & editing, Writing – original draft, Methodology. **Qiuqi Wang:** Writing – review & editing, Writing – original draft, Methodology.

Declaration of competing interest

No potential conflict of interest was reported by the authors.

Acknowledgments

The authors would like to thank the editor-in-chief and anonymous referees for their careful reading and helpful comments on an earlier version of this paper, which leads to a considerable improvement of the presentation of the work. The research of Xia Han is supported by the National Natural Science Foundation of China (Grant Nos. 12301604, 12371471, and 12471449).

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